

# Imaging using ionizing radiations

## *II-SPECT*

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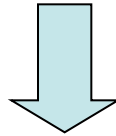
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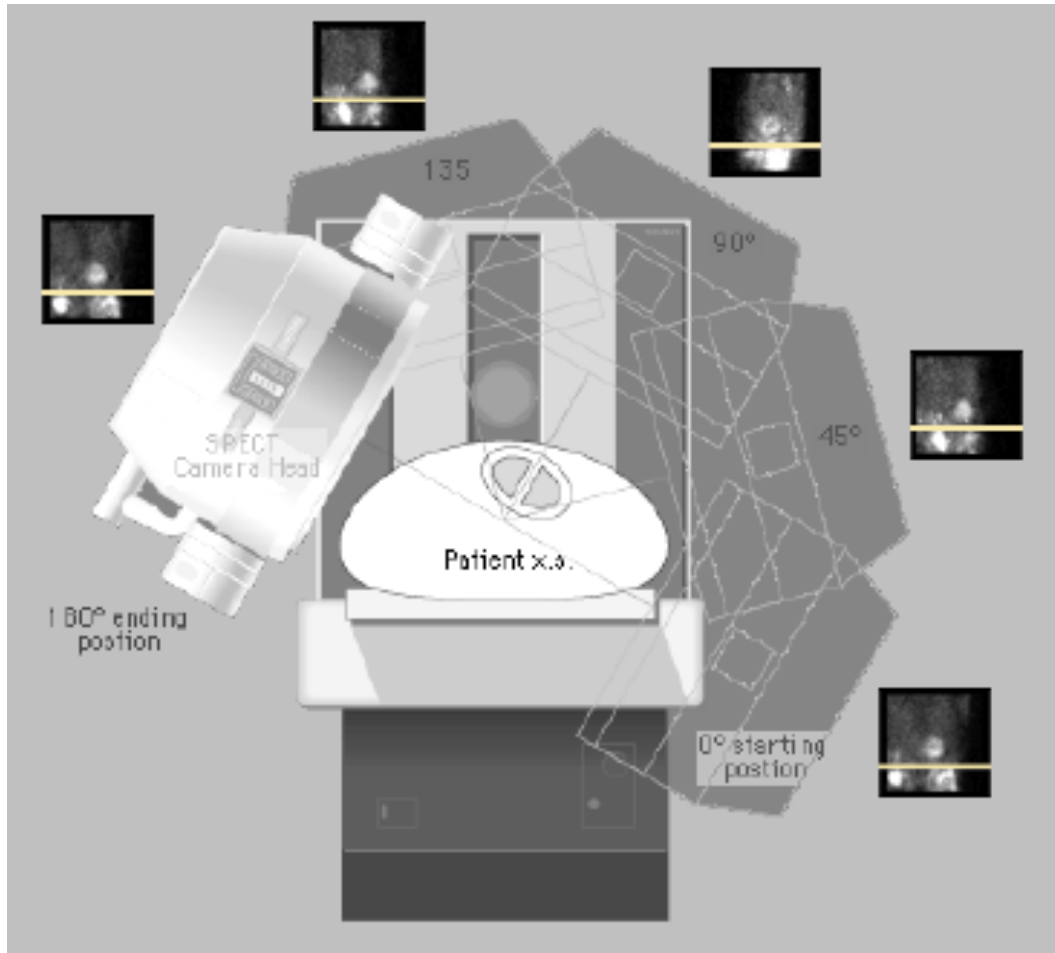
# Tomographic acquisition

- More than a single projection is required in order to obtain the radiotracer distribution.
  - Many possibilities for the solution



- Increasing the number of projections
  - Reduce the number of possibilities

Uniqueness of the solution for an infinity of projections

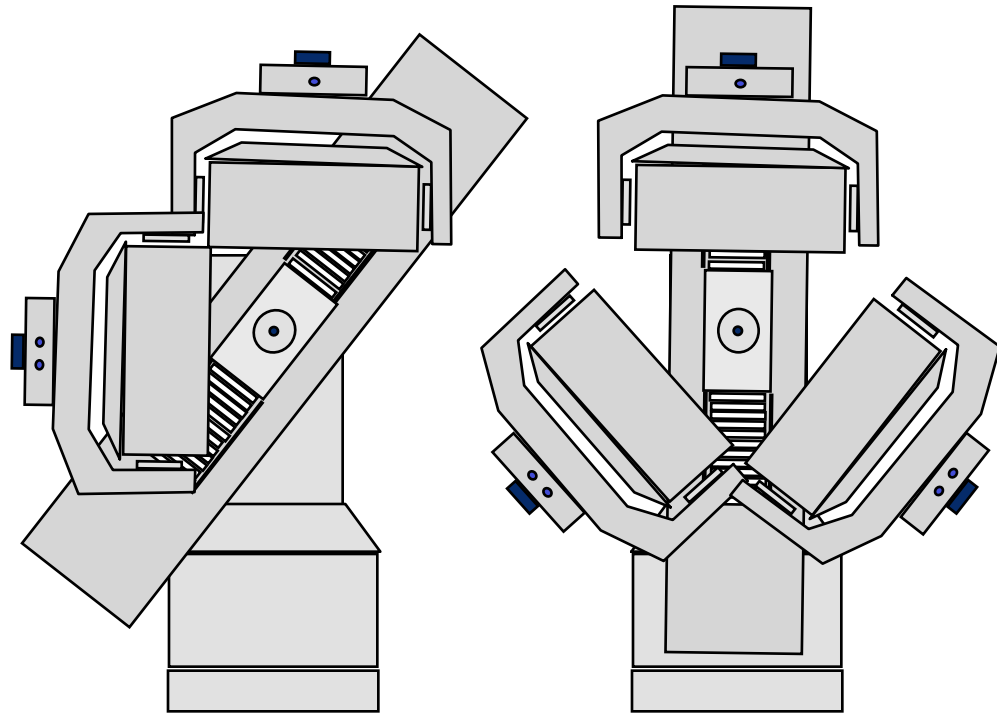
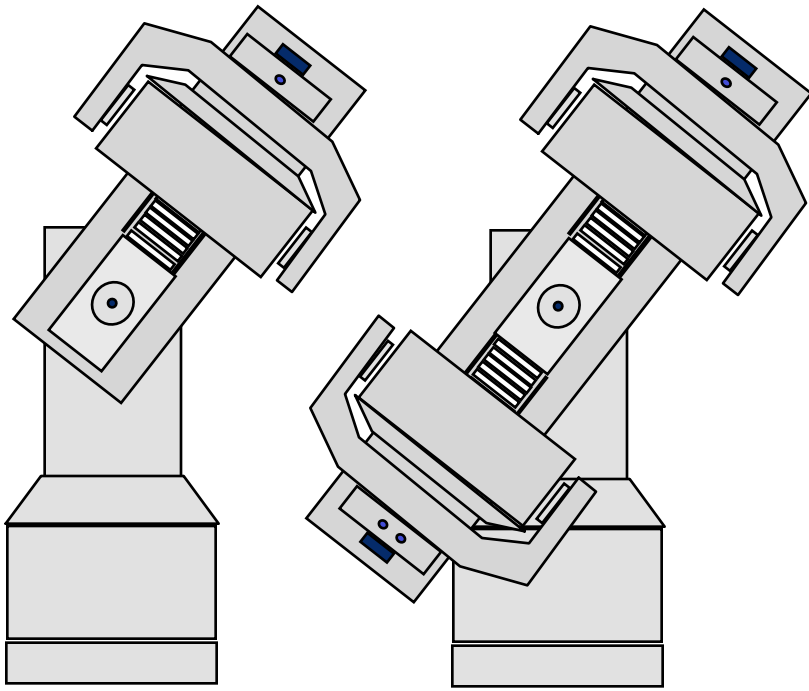


2D projection

Several angles

Reconstruction

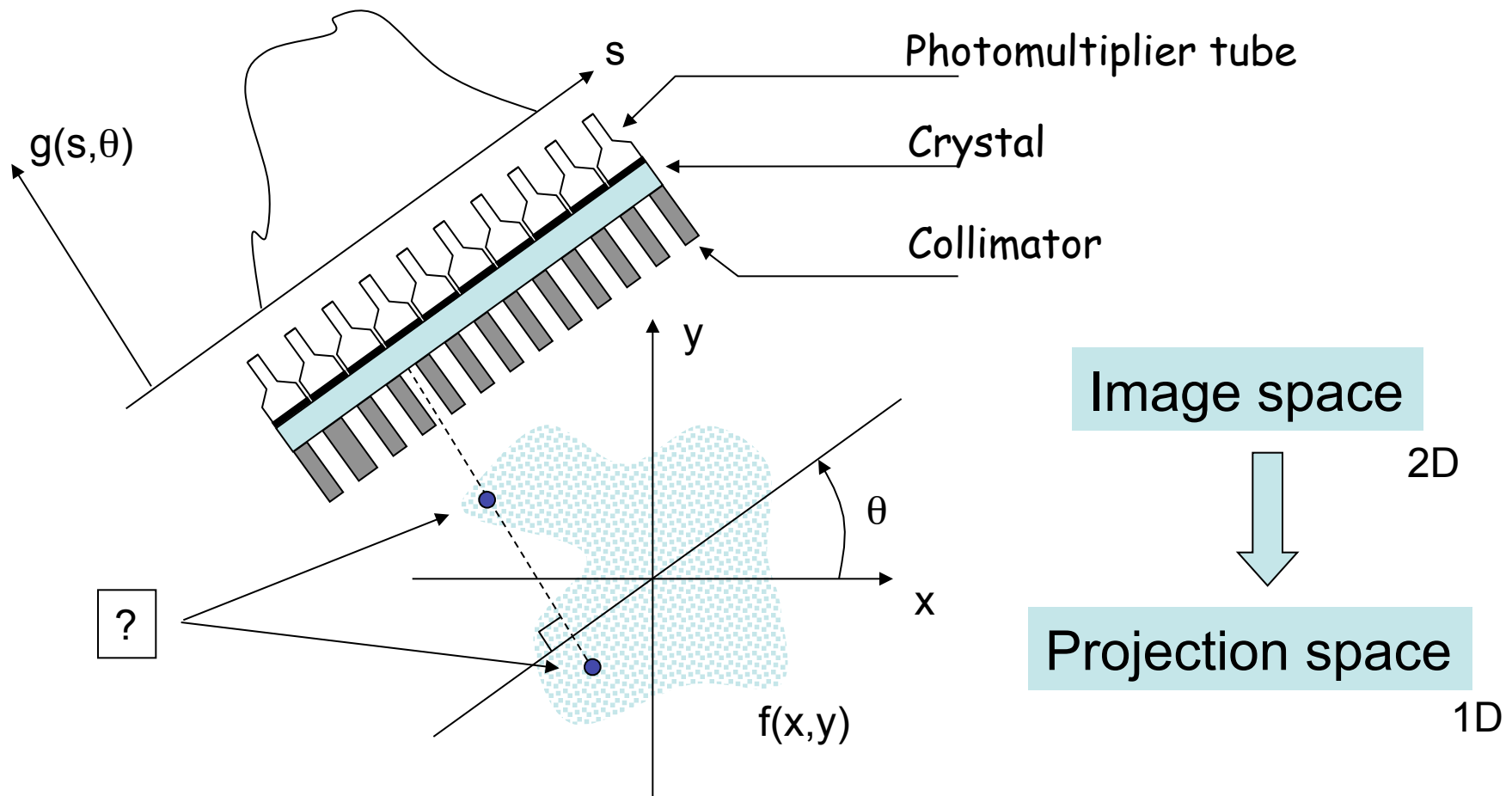
3D Volume



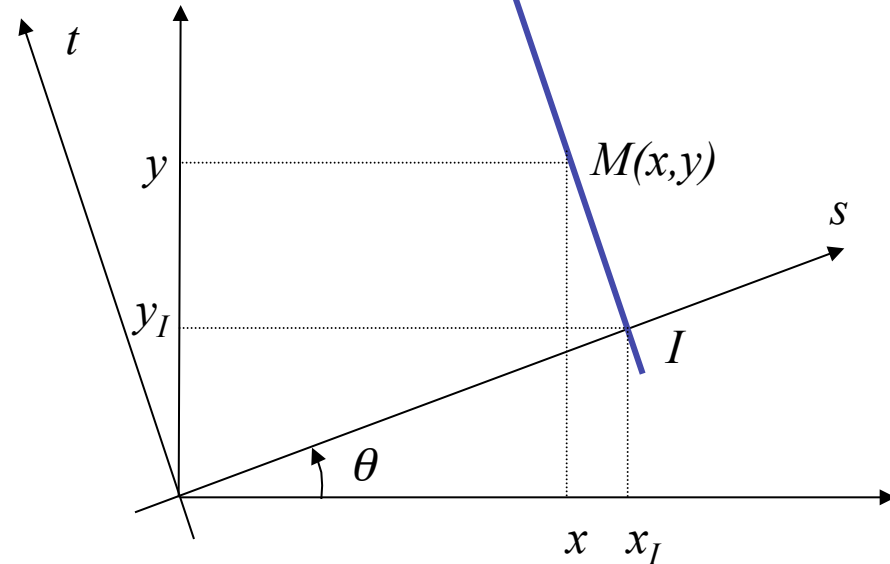
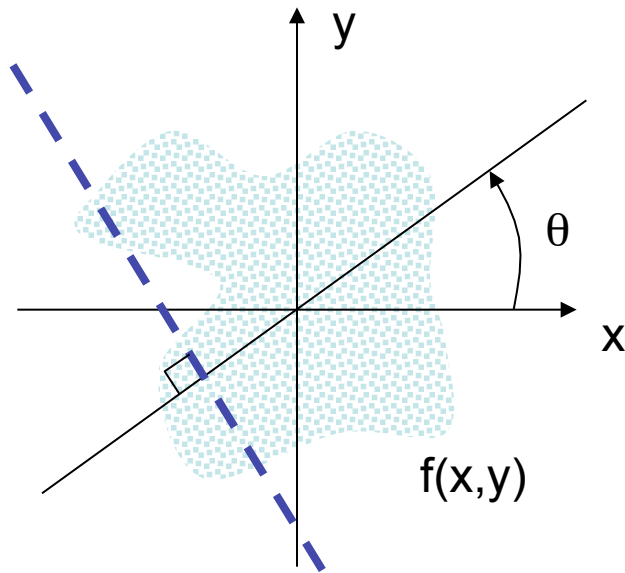
# Positionning the problem

- Emission imaging
  - Injection of a radiopharmaceutical
  - Marker/tracer coupling
- Observables
  - Distribution of  $\gamma$  emitters in the field of view of the camera.
  - $f(x,y)$  : Estimation of the number of photons emitted at  $(x,y)$ .
- Hypothesis:
  - $f(x,y)$  is proportional to the concentration of the injected product.

# Principle of acquisition



# Projection operation



$$\begin{aligned} x_I &= s \cos \theta \\ y_I &= s \sin \theta \end{aligned}$$

$$\begin{aligned} x_I - x &= t \sin \theta \\ y_I - y &= t \cos \theta \end{aligned}$$

$$\begin{aligned} x &= s \cos \theta - t \sin \theta \\ y &= s \sin \theta + t \cos \theta \end{aligned}$$

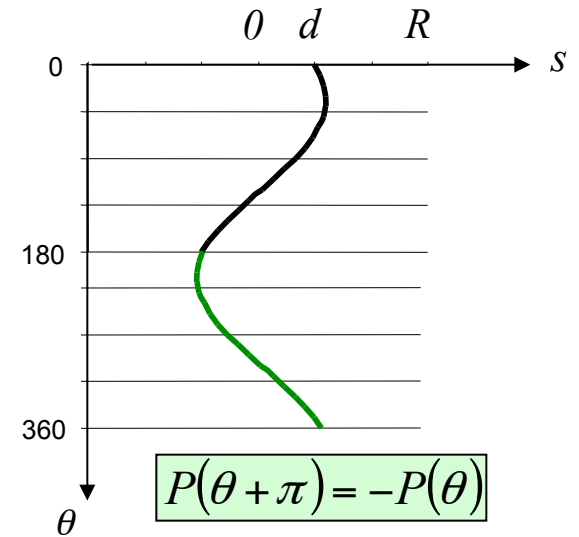
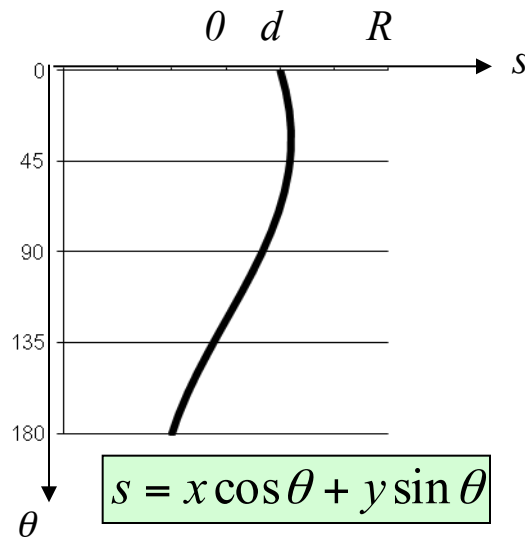
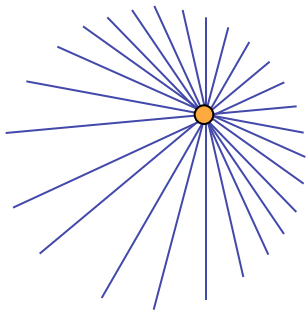
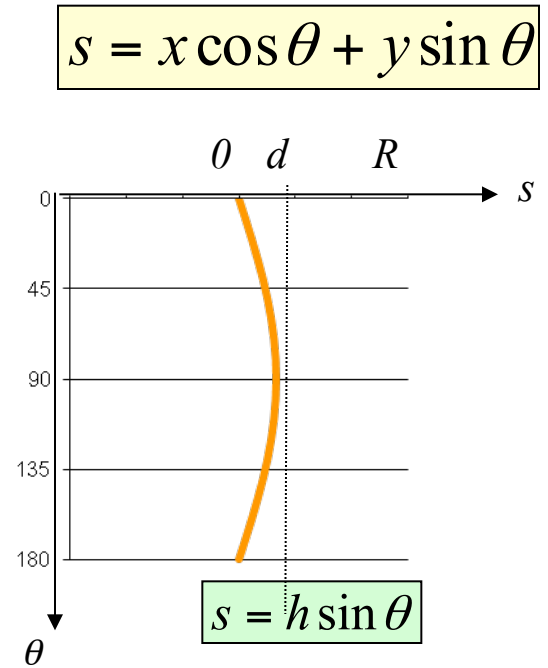
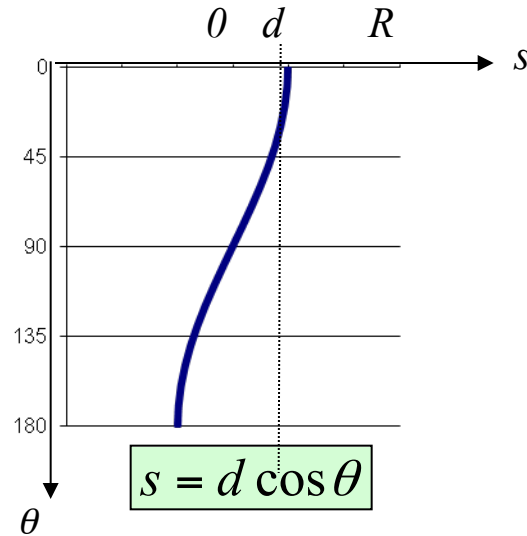
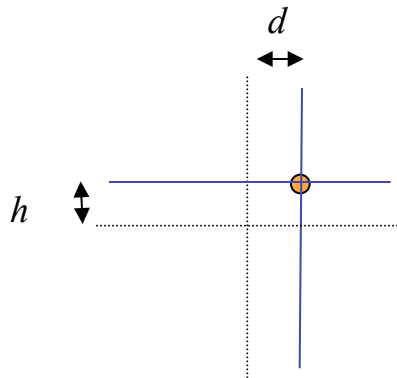
$$\begin{aligned} s &= x \cos \theta + y \sin \theta \\ t &= -x \sin \theta + y \cos \theta \end{aligned}$$



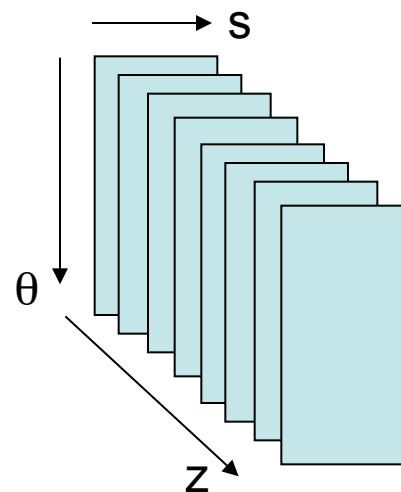
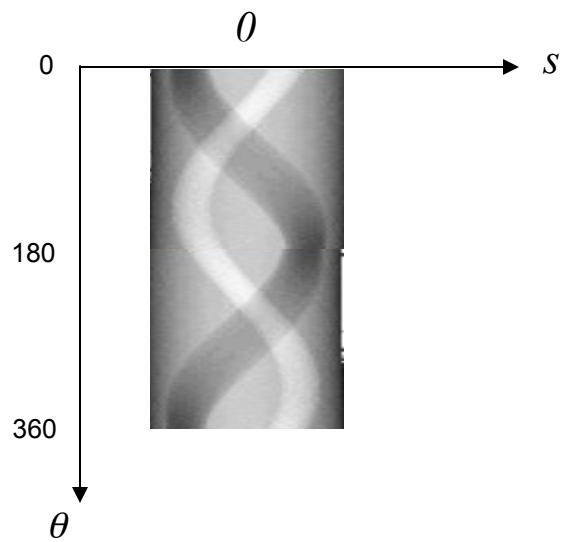
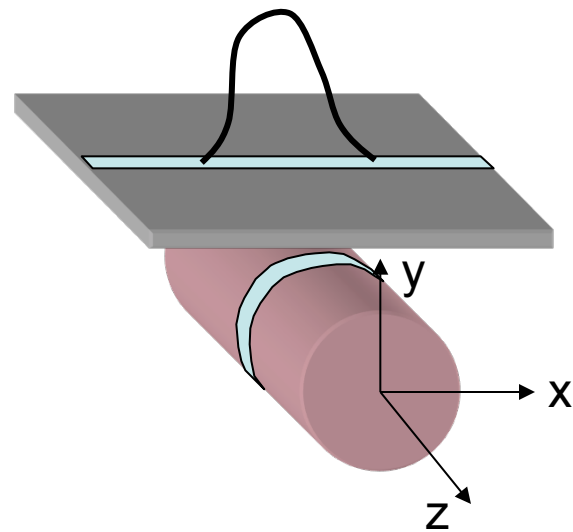
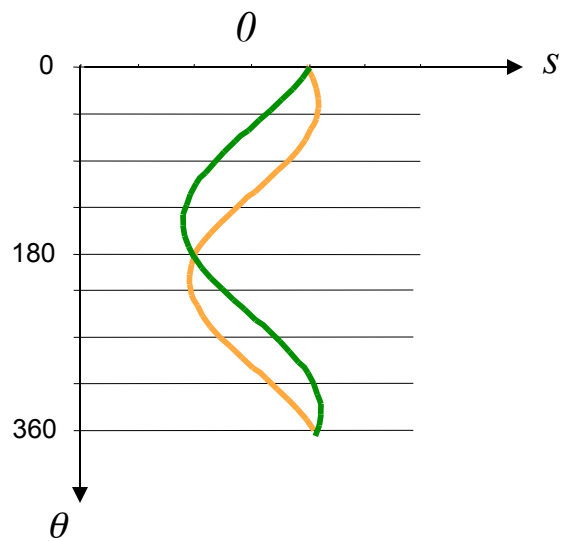
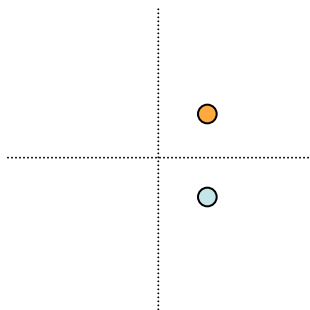
$$s = x \cos \theta + y \sin \theta$$

All points  $M(x,y)$  describing the LOR

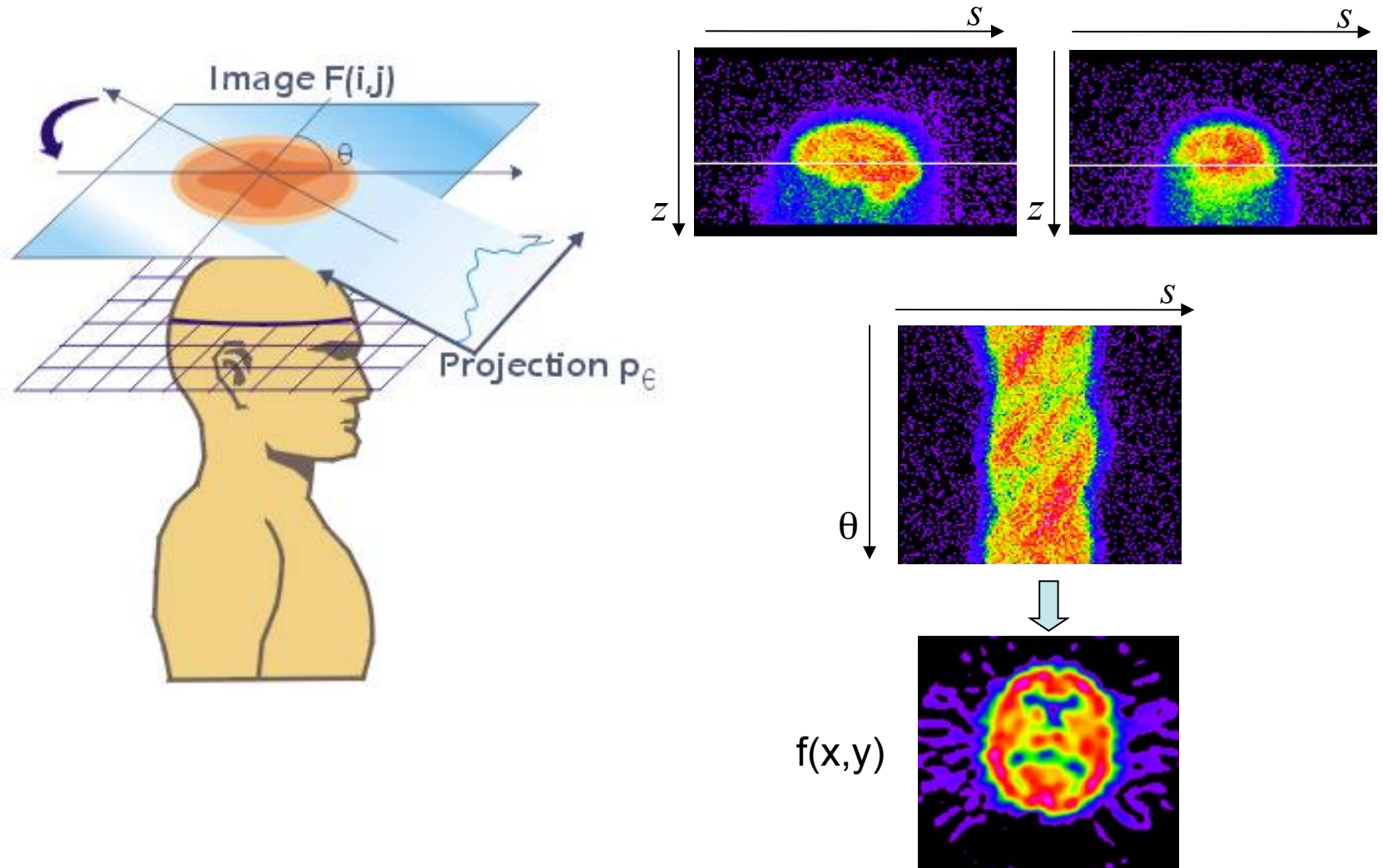
# Sinogramme







# Illustration 2D



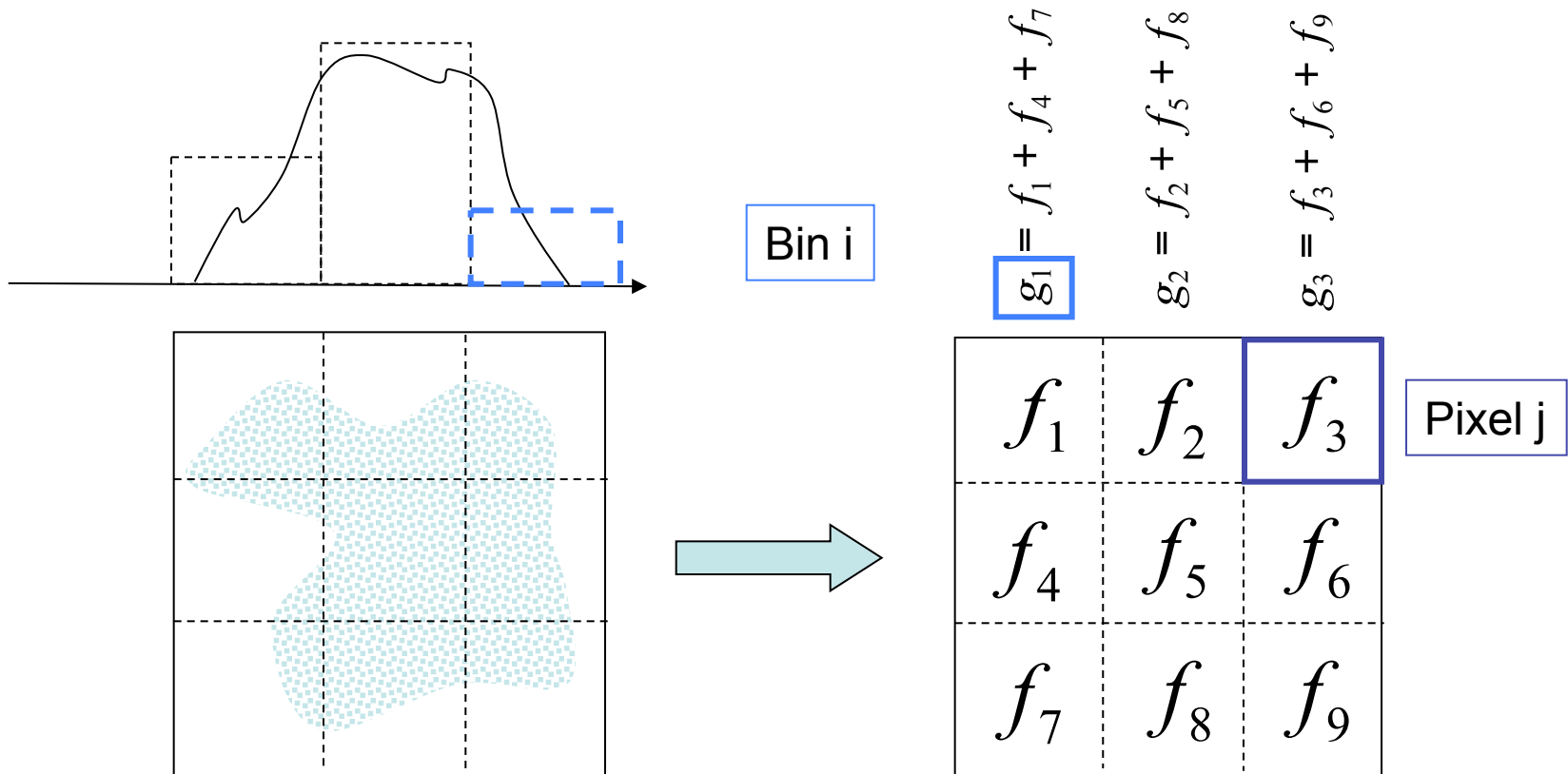
# Radon Transform

In mathematics, the projection operation is defined by the Radon transform

Radon transform  $g(s, \theta) =$   
Integral of  $f(x, y)$  along the line  $D'$

$$g(s, \theta) = \int_{-\infty}^{+\infty} f(s \cos \theta - t \sin \theta, s \sin \theta + t \cos \theta) dt$$

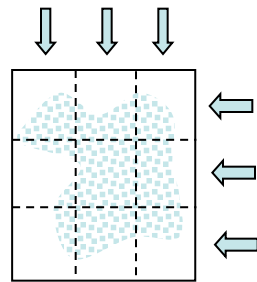
# Continuous to discrete



$$g_i = a_{i1}f_1 + a_{i2}f_2 + \cdots + a_{im}f_m = \sum_{j=1}^m a_{ij}f_j$$

# Matrix representation

$$g = Af$$



$$\begin{bmatrix} g_1 \\ g_2 \\ g_3 \\ g_4 \\ g_5 \\ g_6 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 \\ 0 & 1 & 0 & 0 & 1 & 0 & 0 & 1 & 0 \\ 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 & 1 \\ 1 & 1 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 & 1 & 1 \end{bmatrix} \begin{bmatrix} f_1 \\ f_2 \\ f_3 \\ f_4 \\ f_5 \\ f_6 \\ f_7 \\ f_8 \\ f_9 \end{bmatrix}$$

$$\begin{aligned} g_1 &= f_1 && + f_4 && + f_7 \\ g_2 &= & f_2 && + f_5 && + f_8 \\ g_3 &= & & f_3 && + f_6 && + f_9 \\ g_4 &= f_1 + f_2 + f_3 \\ g_5 &= && f_4 + f_5 + f_6 \\ g_6 &= && && f_7 + f_8 + f_9 \end{aligned}$$

$$\underbrace{\begin{bmatrix} g_1 \\ g_2 \\ g_3 \\ g_4 \\ g_5 \\ g_6 \end{bmatrix}}_g = \underbrace{\begin{bmatrix} 1 & 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 \\ 0 & 1 & 0 & 0 & 1 & 0 & 0 & 1 & 0 \\ 0 & 0 & 1 & 0 & 0 & 1 & 0 & 0 & 1 \\ 1 & 1 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 & 1 & 1 \end{bmatrix}}_A \underbrace{\begin{bmatrix} f_1 \\ f_2 \\ f_3 \\ f_4 \\ f_5 \\ f_6 \\ f_7 \\ f_8 \\ f_9 \end{bmatrix}}_f$$

# Definitions

- *A*: Projection operator
- *a<sub>ij</sub>*: Weight factor representing the contribution of pixel *j* to the number of counts detected in bin *i*.
- *In other words*: probability that a photon emitted from pixel *j* is detected in bin *i*.

# Problem inversion

In theory, direct methods exist to solve the equation:

$$g = Af$$

These methods, called **direct inversion** consist in finding  $A^{-1}$

$$f = A^{-1}g$$

Many difficulties

- Inversion of  $A$
- $A^{-1}$  does not exist
- $A^{-1}$  is not unique

## In practice: inverse problem are badly conceived

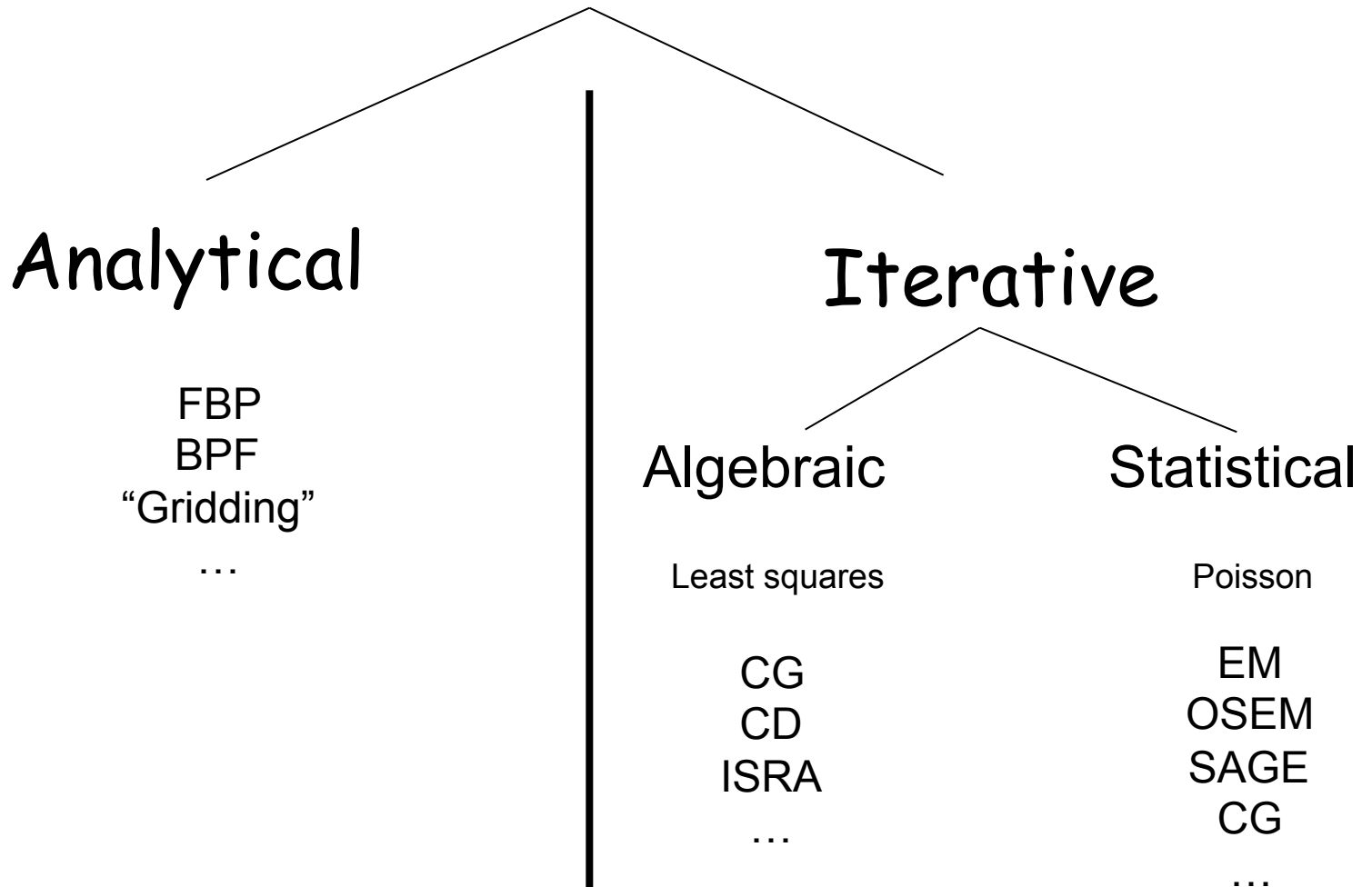
- Solution is not unique and  $A$  is unstable:
  - Data contamination by noise
  - Finite number of projections
- Approached solution

### Problem:

Knowing the sinogramm,  
What is the radioactive distribution  $f(x,y)$ ?



# Image reconstruction



# Analytic algorithms

- Backprojection operation
- Central slice theorem
- Backprojection + filtering
- Backprojection of filtered projections
- Fourier domain (space)

# Operator: Back Projection

Inverse operator

$$b(x, y) = \int_0^{\pi} g(s, \theta) d\theta$$

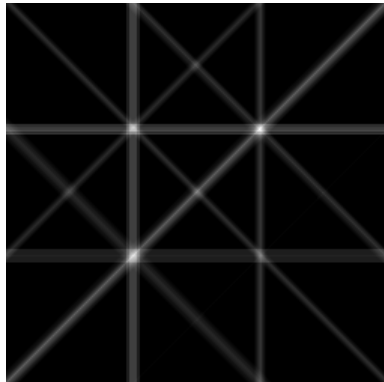
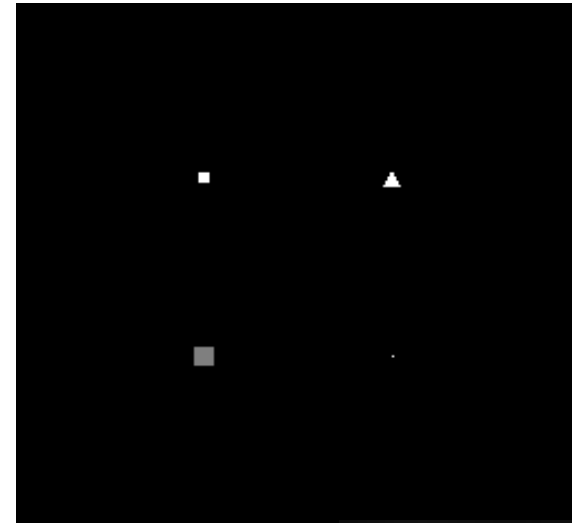
$$\tilde{b}(x, y) = \sum_{k=1}^p g(s_k, \theta_k) \Delta\theta$$

$p$ : number of projections

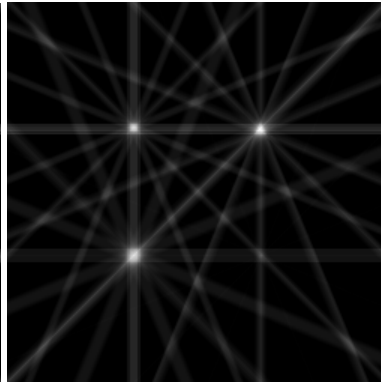
$\Delta\theta$ : Sampling ( $\pi/p$ )

# Back projection: Artifacts

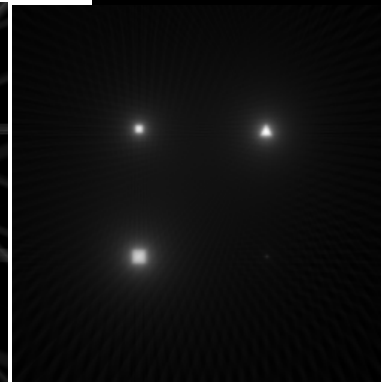
$g_1$	$g_2$	$g_3$	
$\downarrow$	$\downarrow$	$\downarrow$	
$\frac{g_1 + g_4}{2}$	$\frac{g_2 + g_4}{2}$	$\frac{g_3 + g_4}{2}$	$\leftarrow g_4$
$\frac{g_1 + g_5}{2}$	$\frac{g_2 + g_5}{2}$	$\frac{g_3 + g_5}{2}$	$\leftarrow g_5$
$\frac{g_1 + g_6}{2}$	$\frac{g_2 + g_6}{2}$	$\frac{g_3 + g_6}{2}$	$\leftarrow g_6$



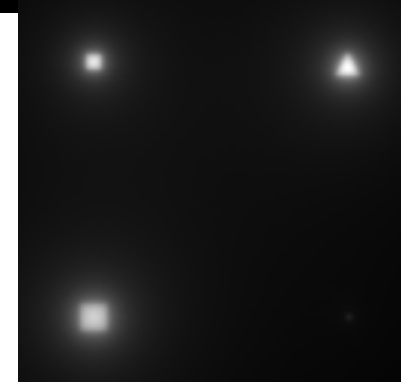
$p=4$



$p=8$



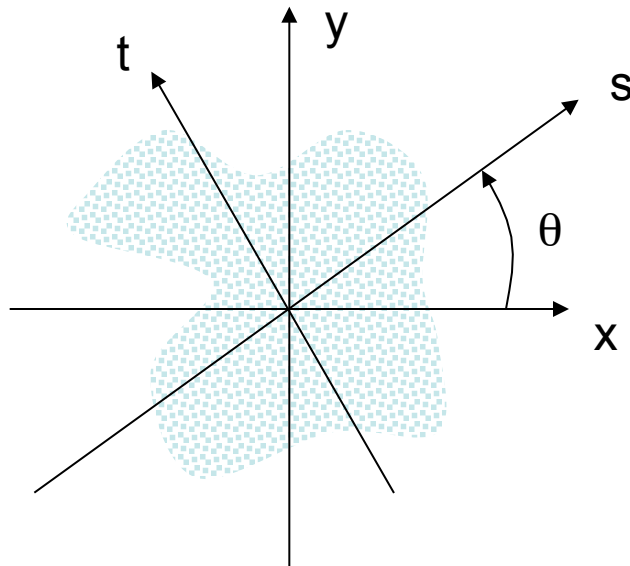
$p=64$



$p=256$

# Central slice theorem

$$g(s, \theta) = \int_{-\infty}^{+\infty} f(x, y) dt$$



TF (g)

$$G_{10}(v_s, \theta) = \int_{-\infty}^{+\infty} g(s, \theta) e^{-2i\pi v_s s} ds$$

$$G_{10}(v_s, \theta) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x, y) e^{-2i\pi v_s s} ds dt$$

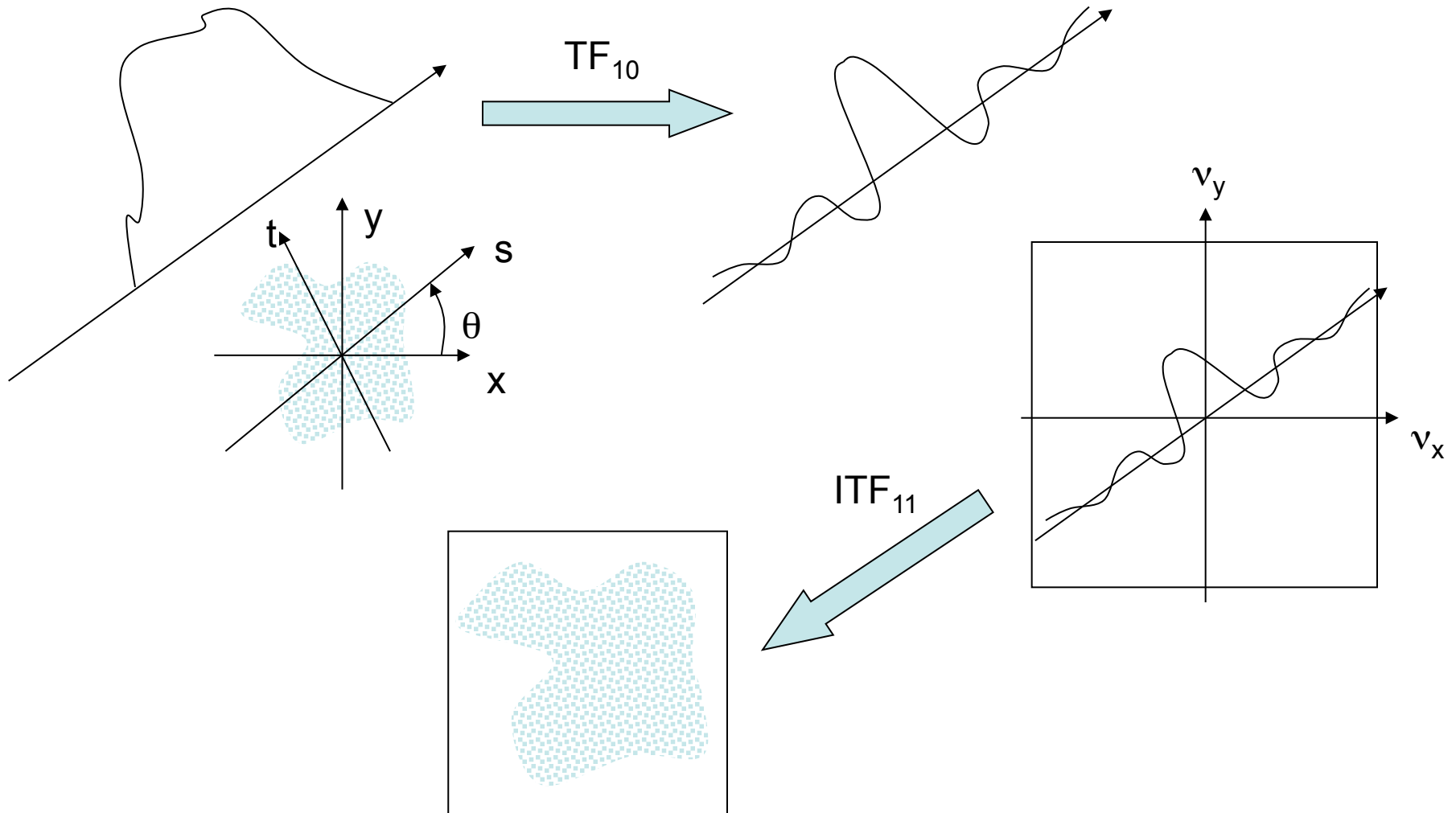
$$s = x \cos \theta + y \sin \theta$$

$$\begin{aligned} v_x &= v_s \cos \theta \\ v_y &= v_s \sin \theta \end{aligned}$$

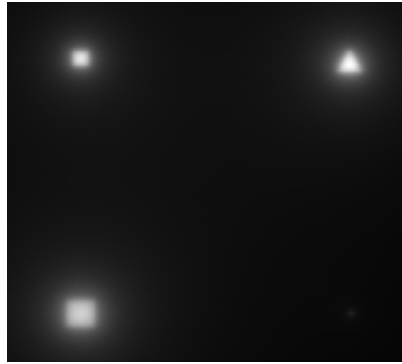
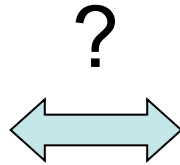
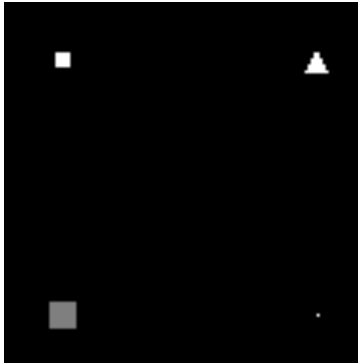
$$G_{10}(v_s, \theta) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x, y) e^{-2i\pi(xv_x + yv_y)} dx dy$$

$$F_{11}(v_x, v_y) \Big|_{v_t=0} = G_{10}(v_s, \theta)$$

# Graphical illustration



# Sampling

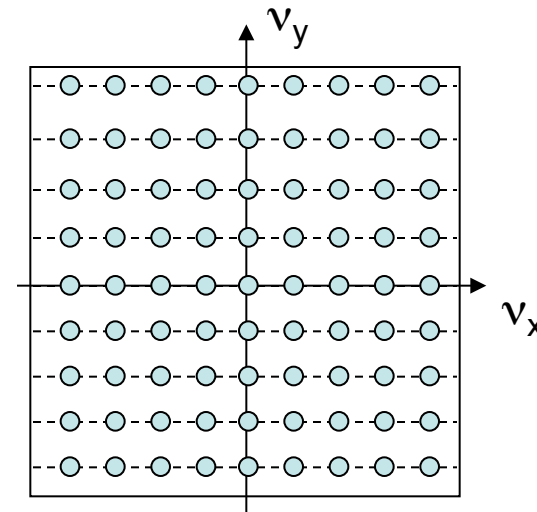
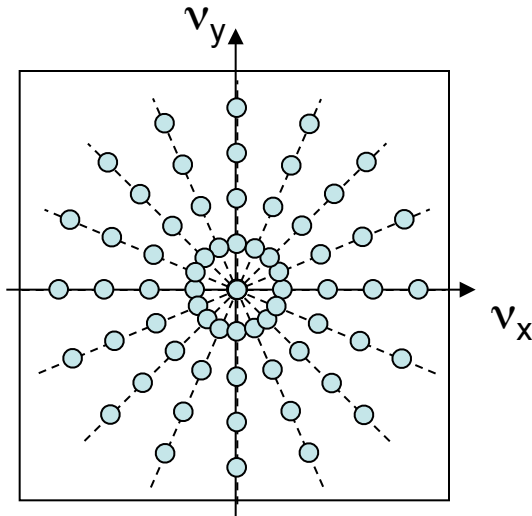


Low frequencies  
upsampling



Raising high  
frequencies

Polar -> Cartesien



# Proof

$$f(x, y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} F_{11}(v_x, v_y) e^{2i\pi(xv_x + yv_y)} dv_x dv_y$$

TF 2D

$$f(x, y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} G_{10}(v_s, \theta) e^{2i\pi(xv_x + yv_y)} dv_x dv_y$$

$$F_{11}(v_x, v_y) = G_{10}(v_s, \theta)$$

Changement de variables:

$$(v_x, v_y) \rightarrow (v_s, \theta)$$

$$v_x = v_s \cos \theta$$

$$v_y = v_s \sin \theta$$

$$s = x \cos \theta + y \sin \theta$$

$$f(x, y) = \int_0^{\pi} \int_{-\infty}^{+\infty} G_{10}(v_s, \theta) |v_s| e^{2i\pi v_s s} dv_s d\theta$$

$$dv_x dv_y = |v_s| dv_s d\theta$$

$$f(x, y) = \int_0^{\pi} g'(s, \theta) d\theta$$



# Filtering

- Exact inversion is not possible for two reasons:
  - Discrete sampling -> Limited space
    - Shannon: fréquence max. reconstructed :  
 $Nyquist = 1/2\Delta s$
  - Presence of statistical noise
    - Utilization of « ramp » filter -> Noise amplification



Utilisation of an apodisation window

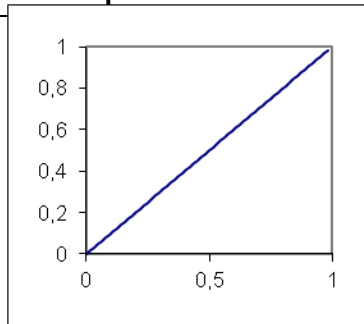
# Apodisation window

- Cut-off frequency influence:
  - The resolution of the reconstructed image
  - Noise properties

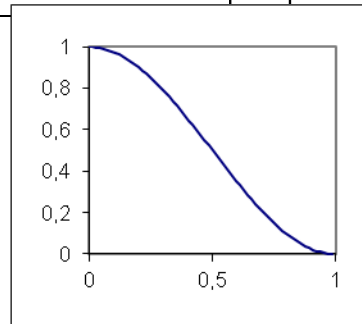
## Hann Filter

$$W(v_s) = \begin{cases} 0,5 \left( 1 + \cos \left( \frac{\pi v_s}{v_c} \right) \right), & |v_s| < v_c \\ 0 & |v_s| \geq v_c \end{cases}$$

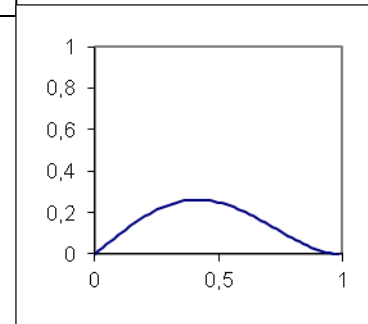
$$0 \leq \alpha \leq 1$$



$|v_s|$



$W(v_s)$



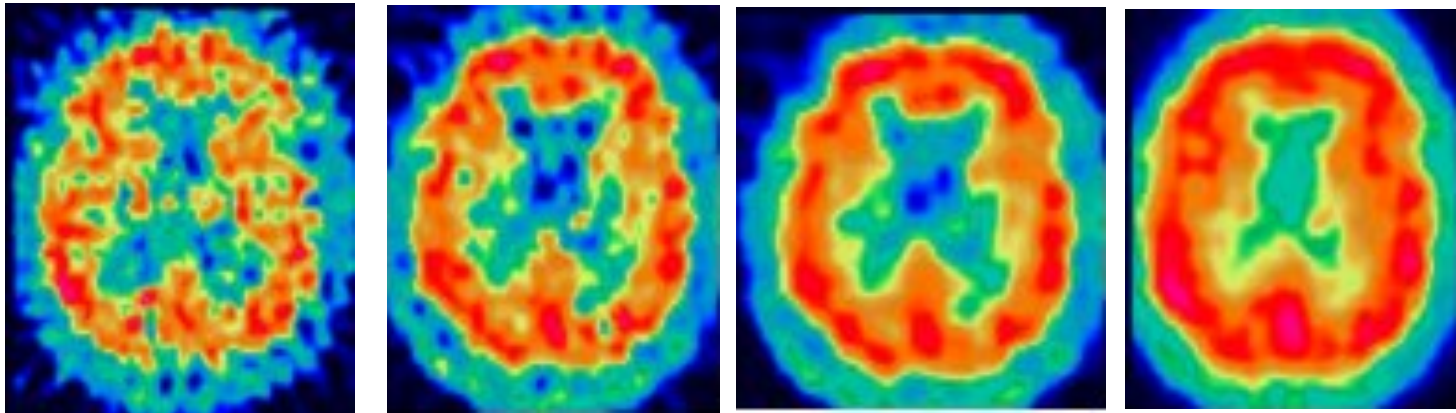
$|v_s| \times W(v_s)$

# Cut-off frequency: Resolution



←  $\nu_c$

# Cut-off frequency: Noise

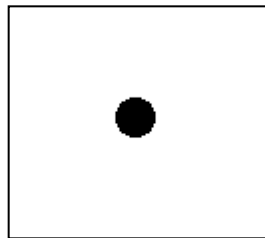


$\nu_c$

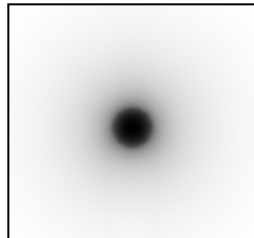
# Back projection + Filtering

$$f(x, y) = b(x, y) \otimes \text{psf}(x, y)$$

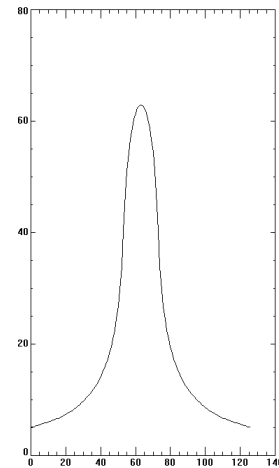
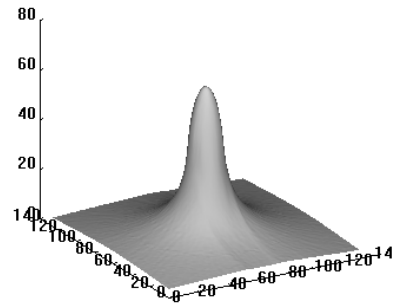
- Drawback: requires an entire huge back projection matrix  $b(x, y)$ 
  - Otherwise, truncation artifacts



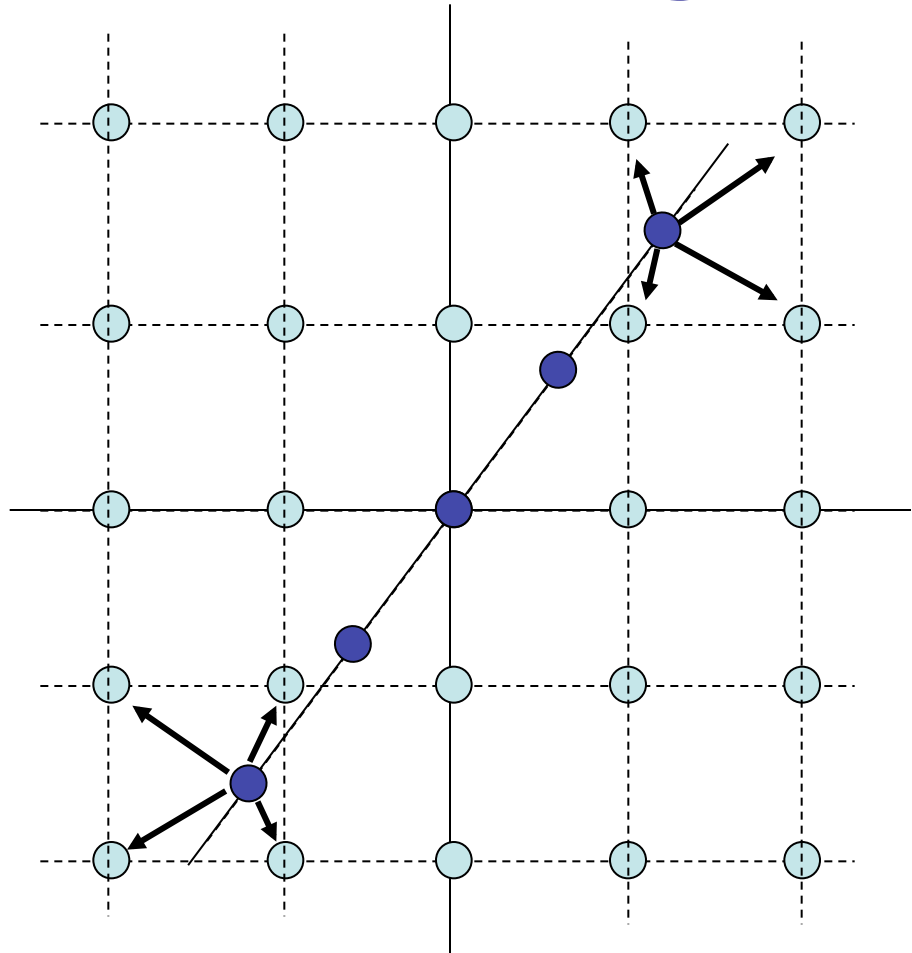
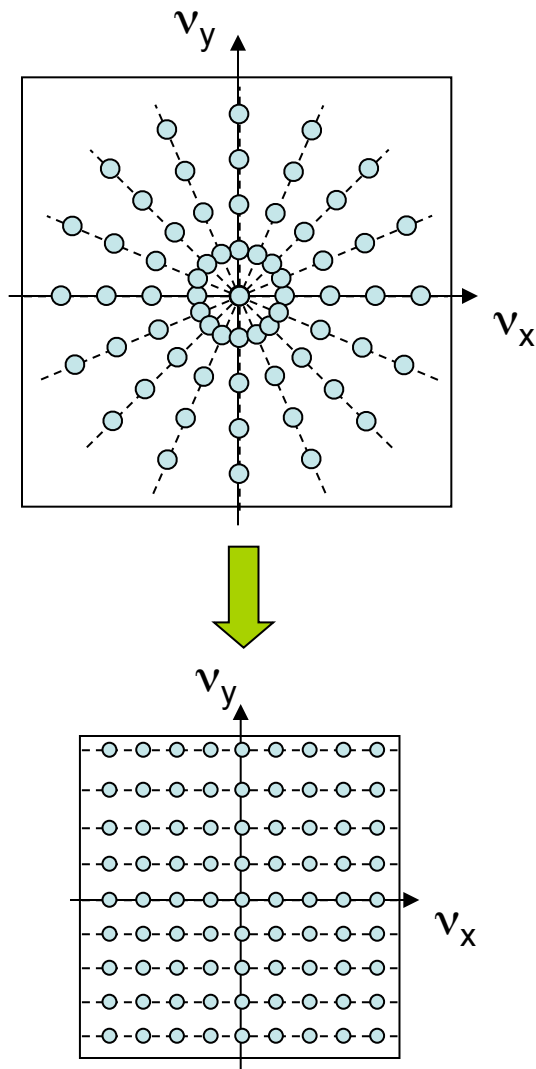
$f(x, y)$



$b(x, y)$

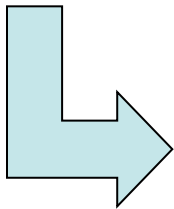


# Fourier space: Gridding

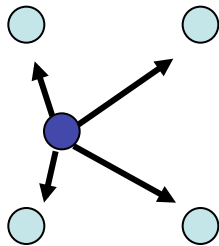


# Drawback of fourier methods

- Sensitivity to noise
- Interpolation kernel in Fourier



Multiplication in real space



$$F_{11}(v_x, v_y) \otimes W(v_x, v_y) \rightarrow f(x, y) \cdot w(x, y)$$

# Iterative algorithms

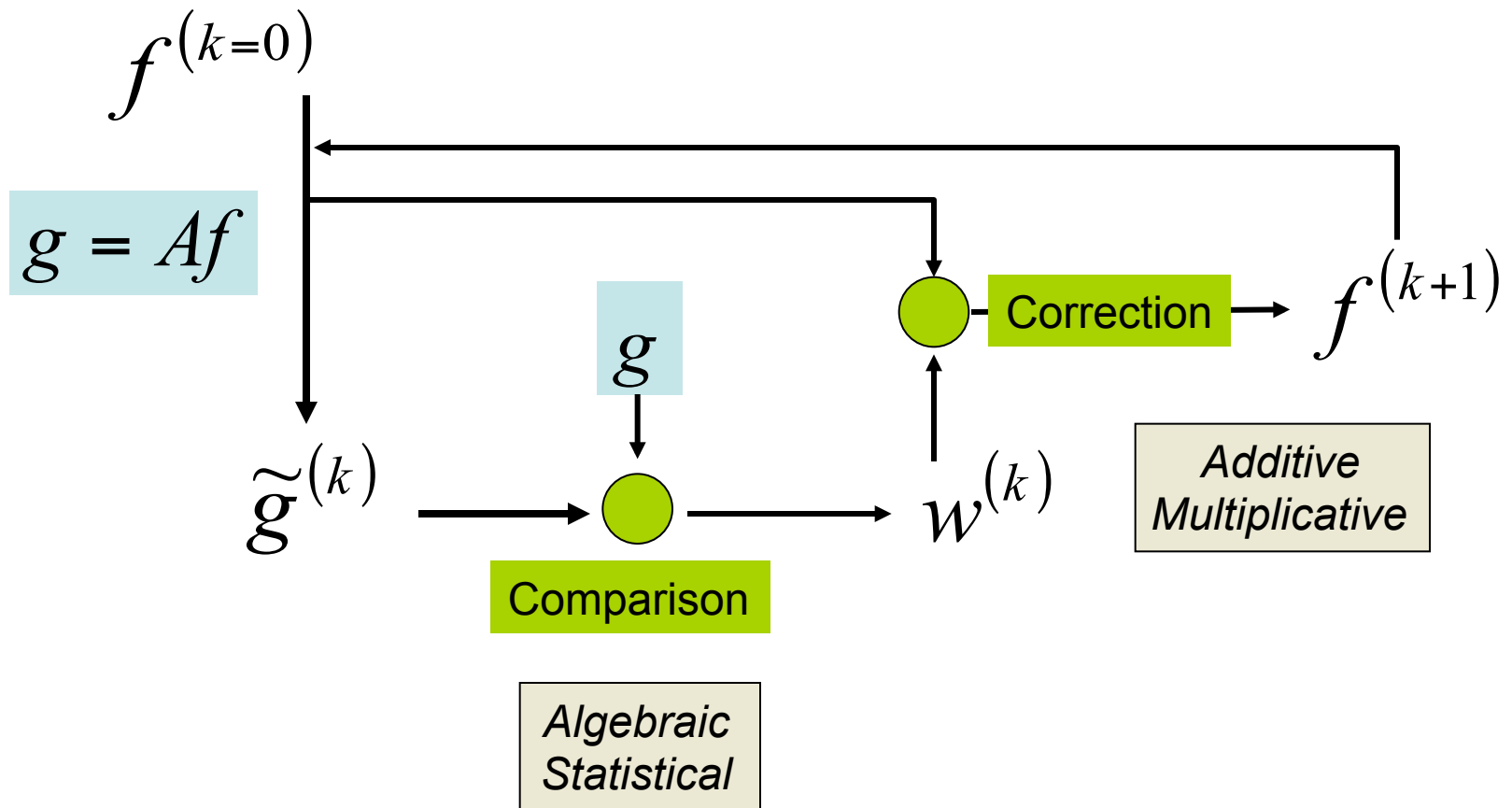
- Find the  $f$  vector, solution of the equation

$$g = Af$$

- Iterative algorithms are based on the principle of finding a solution by successive estimations.
- The projection corresponding to the current estimation is compared to the acquired projections.
- The comparison result is used to modify the estimation and create a new one.



# Principle



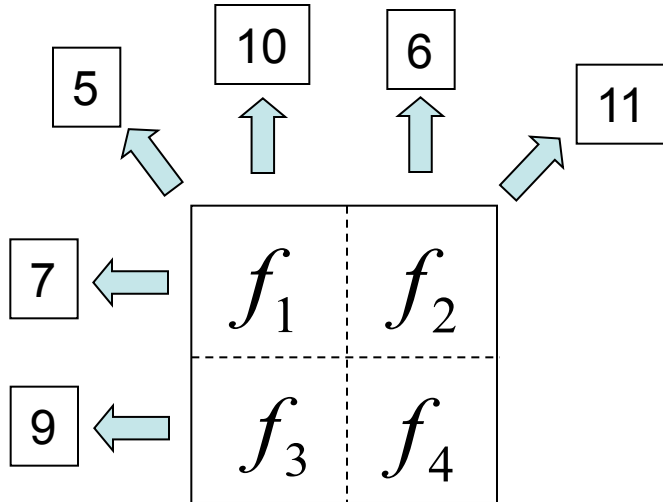
# Algebraic methods: ART

- ART: « Algebraic Reconstruction Technique »

$$f_j^{(k+1)} = f_j^{(k)} + \frac{g_i - \sum_{j=1}^N f_{ji}^{(k)}}{N}$$

# Example ART-1

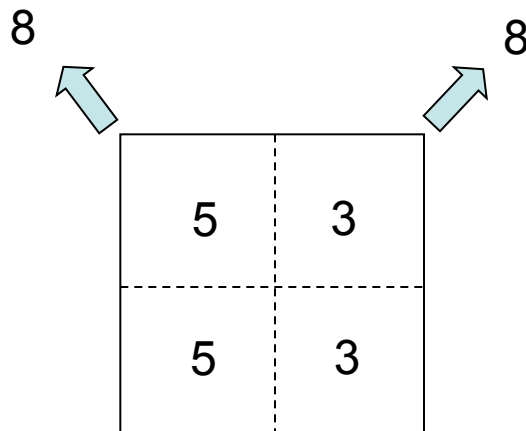
$$f_j^{(k+1)} = f_j^{(k)} + \frac{g_i - \sum_{j=1}^N f_{ji}^{(k)}}{N}$$



0	0
0	0

$$f_1^{(1)} = f_3^{(1)} = 0 + \frac{10 - 0}{2} = 5$$

$$f_2^{(1)} = f_4^{(1)} = 0 + \frac{6 - 0}{2} = 3$$



$$f_1^{(2)} = 5 + \frac{5 - 8}{2} = 3,5$$

$$f_2^{(2)} = 3 + \frac{11 - 8}{2} = 4,5$$

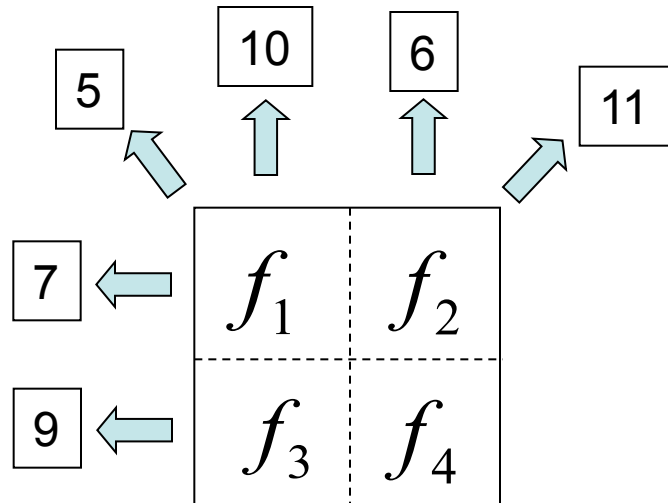
$$f_3^{(2)} = 5 + \frac{11 - 8}{2} = 6,5$$

$$f_4^{(2)} = 3 + \frac{5 - 8}{2} = 1,5$$

3,5	4,5
6,5	1,5

# Example ART-2

$$f_j^{(k+1)} = f_j^{(k)} + \frac{g_i - \sum_{j=1}^N f_{ji}^{(k)}}{N}$$



3	4
7	2

8	3,5	4,5
8	6,5	1,5

$$f_1^{(3)} = 3,5 + \frac{7-8}{2} = 3$$

$$f_2^{(3)} = 4,5 + \frac{7-8}{2} = 4$$

$$f_3^{(3)} = 6,5 + \frac{9-8}{2} = 7$$

$$f_4^{(3)} = 1,5 + \frac{9-8}{2} = 2$$

# Why using statistical methods?

## Advantages:

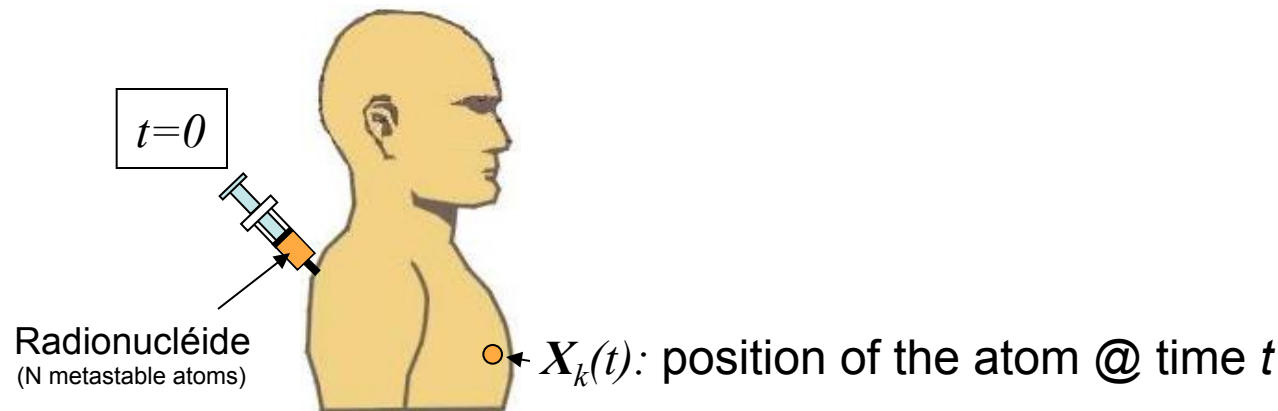
- Constraints on the object: non negativity, support...
- Incorporation of physics models
  - Photons transportation / geometrical efficiency...
- Appropriated statistical model (Noise reducing).
- Flexibility regarding geometry.
- Incorporation of anatomical information.

## Drawbackss:

- Computation time.
- Complex model.
- Tedious implementation.

# Investigating the object

- Realize the image of the radiotracer distribution



Imaging system: provide  $X_k(t)$  ?

First hypothesis:  $X_k(t)$  Independent random variables distributed according to the same probability density function  $f_{\vec{X}(t)}(\vec{x})$

Imaging system: provide  $f_{\vec{X}(t)}(\vec{x})$  !!

Secone hypothesis: Atoms distribution follows a Poisson law

# Radioactive decay

An atom can only be observed when it deexcitates and emits photons.  
The deexcitation time of an atom  $k$ th is a random variable  $T_k$ .

Third hypothesis: The  $T_k$  are independent random variables

Fourth hypothesis: Each  $T_k$  has an exponential distribution whose mean is  $\mu_T = t_{1/2}/\ln 2$

$$t_{1/2} = \text{« half-life »}$$

The photon emission is a statistical process that follows a Poisson law

# Statistics of an ideal counter

$K(t, V)$ : number of atoms @ time  $t$ , located in a volume  $V$

$K(t, V)$ : counting process following a Poisson law with a mean

$$E[K(t, V)] = \int_0^t \int_V \lambda(\vec{x}, \tau) d\vec{x} d\tau$$

with

$$\lambda(\vec{x}, t) = \mu_N \frac{e^{-t/\mu_T}}{\mu_T} \cdot f_{\vec{X}(t)}(\vec{x})$$



# Detection element

For example: one element of the sinogram (bin)



does not correspond necessarily to a physical element of the detector

Fifth hypothesis: Each desintegration produce a detected event in one bin at least.

If a fraction of the event is attributed to 1 bin  
⇒ Counting statistics follows a law different than Poisson.

# Detection efficiency

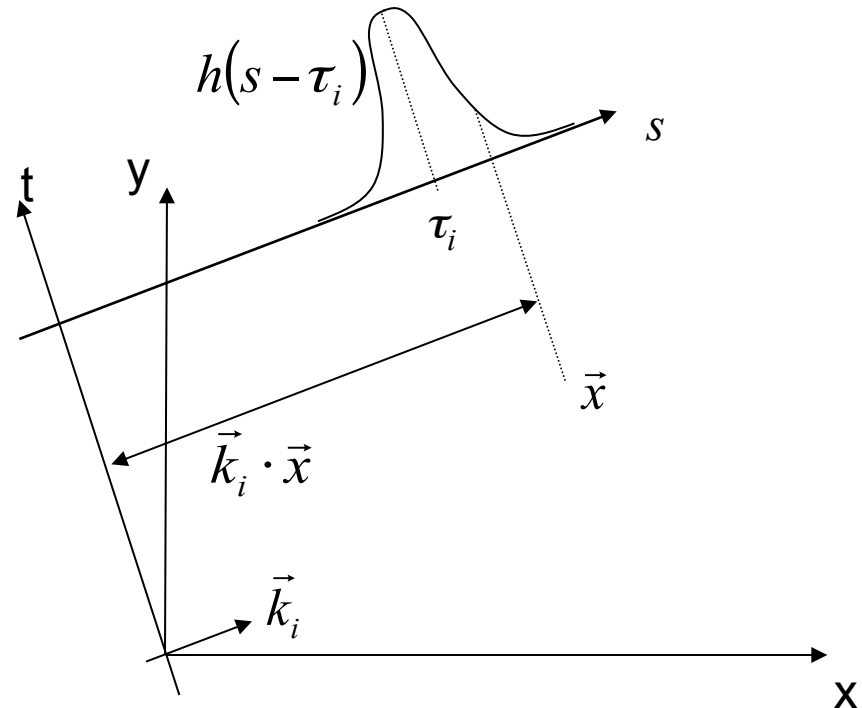
$s_i(\vec{x})$  Probability of detecting in bin  $i$ , an event coming from position  $x$

PSF: Impulse response of the detection system  
« Point Spread Function »

$$s_i(\vec{x}) = h(\vec{k}_i \cdot \vec{x} - \tau_i)$$

$$h(s) = \delta(s)$$

Ideal detector



# Detection efficiency

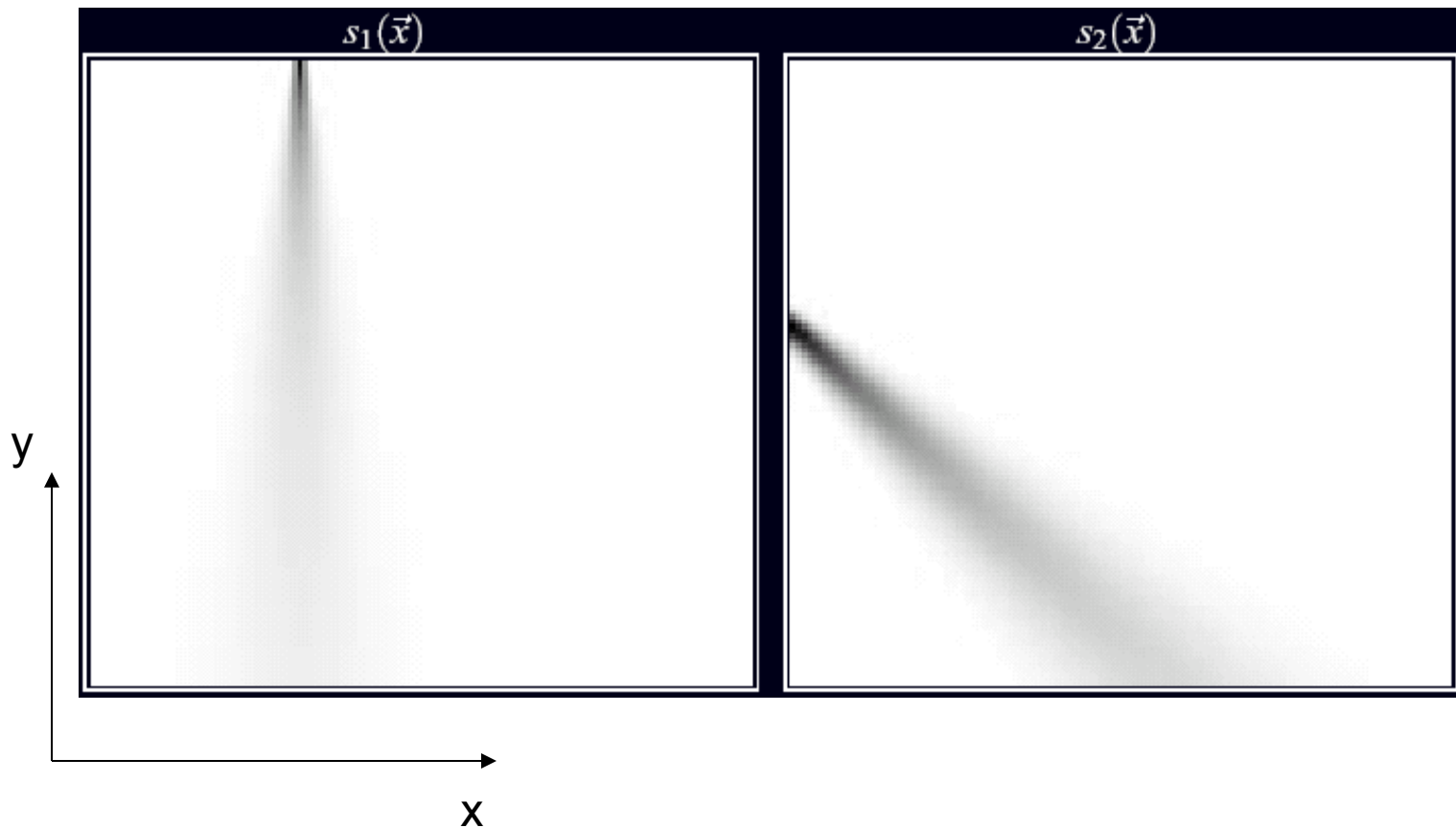
$$s_i(\vec{x})$$

Including:

- The geometry / solid angle of detection
- The collimation
- The scatter
- The attenuation
- Detector response
- Detection efficiency of the detector
- Positron range, acolinearity, etc...

# Examples

Detection efficiency for an Anger gamma camera



# Acquisition

- Register events
  - for  $t$  between  $t_1$  and  $t_2$
- $Y_i$ : number of events registered by the  $i$ th detector element
- $\{Y_i: i=1, \dots, n_d\}$  represents the sinogram data.

In summary,

$$Y_i \sim \text{Poisson} \left\{ \int s_i(\vec{x}) \lambda(\vec{x}) d\vec{x} \right\}$$

With

$$\lambda(\vec{x}, t) = \mu_N \frac{e^{-t/\mu_T}}{\mu_T} \cdot f_{\vec{X}(t)}(\vec{x})$$

$$\lambda(\vec{x}) = \mu_N \int_{t_1}^{t_2} f_{\vec{X}(t)}(\vec{x}) \frac{e^{-t/\mu_T}}{\mu_T} dt \quad = \text{Emission density}$$

# Poisson Statistical Model

Measurements = real events + noise

Sources of noise:

- cosmic rays
- ambient (surrounding) noise
- All counts not taken into account in  $s_i(x)$

$$Y_i \sim \text{Poisson} \left\{ \int s_i(\vec{x}) \lambda(\vec{x}) d\vec{x} + r_i \right\} \quad i = 1, \dots, n_d$$



Mean number of events originated from a noise source in bin  $i$

# Problem posed by reconstruction

Estimate the emission density  $\lambda$  using:

$$Y_i \sim \text{Poisson} \left\{ \int s_i(\vec{x}) \lambda(\vec{x}) d\vec{x} + r_i \right\}, \quad i = 1, \dots, n_d$$

$$\{Y_i = y_i\}_{i=1}^{n_d}$$

Events collected in bin  $i$

$$s_i(\vec{x})$$

Detection efficiency in bin  $i$

$$r_i$$

Noise source in bin  $i$

# In summary: five multiple choice parts

- -1- Object description  $\lambda(\vec{x})$
- -2- Physical model of the system  $s_i(\vec{x})$
- -3- Statistical model of the measurement  $Y_i$
- -4- Optimization criteria
- -5- Used algorithm



# -1- Object description

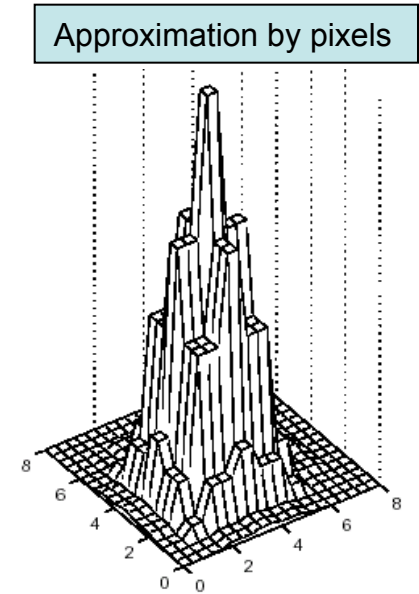
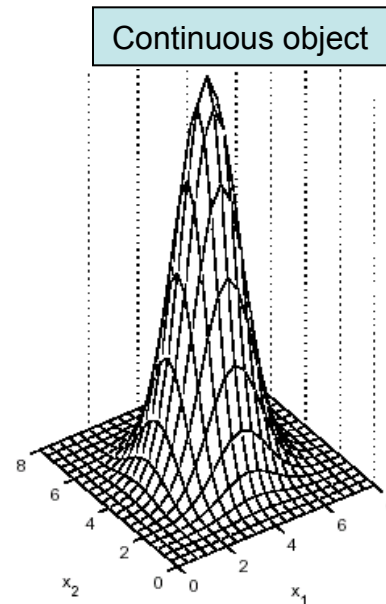
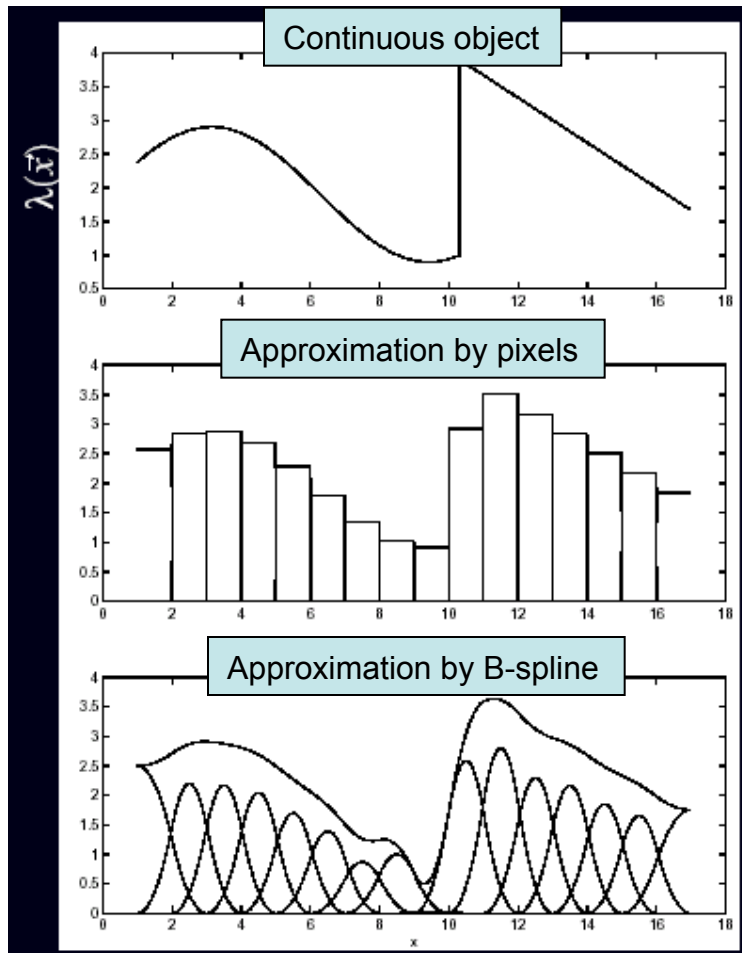
$\lambda(\vec{x})$  is a continuous function

↳ Replaced by  $\lambda = (\lambda_1, \dots, \lambda_{n_p})$

With  $\lambda(\vec{x}) \approx \sum_{j=1}^{n_p} \lambda_j \boxed{b_j(\vec{x})}$  basis function

- Fourier series
- Wavelette
- Kaiser-Bessel
- B-splines
- Rectangular pixels
- Basis on the organs
- ...

# Examples



# -1- Projection algorithm

$$g(s, \theta) = \int_{-\infty}^{+\infty} f(x, y) dt$$

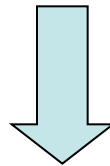
$$g_i = a_{i1}f_1 + a_{i2}f_2 + \cdots + a_{im}f_m = \sum_{j=1}^m a_{ij}f_j$$

$$g(s, \theta) = \int s_i(\vec{x}) \lambda(\vec{x}) d\vec{x} = \int s_i(\vec{x}) \left[ \sum_{j=1}^{n_p} \lambda_j b_j(\vec{x}) \right] d\vec{x}$$

$$= \sum_{j=1}^{n_p} \left[ \int s_i(\vec{x}) b_j(\vec{x}) d\vec{x} \right] \lambda_j = \sum_{j=1}^{n_p} a_{ij} \lambda_j$$

# -1- Discrete Reconstruction

$$Y_i \sim \text{Poisson} \left\{ \int s_i(\vec{x}) \lambda(\vec{x}) d\vec{x} + r_i \right\}, \quad i = 1, \dots, n_d$$



$$Y_i \sim \text{Poisson} \left\{ \sum_{j=1}^{n_p} a_{ij} \lambda_j + r_i \right\}, \quad i = 1, \dots, n_d$$

## -2- Physical model of the Système

$$a_{ij} = \int s_i(\vec{x}) b_j(\vec{x}) d\vec{x}$$

- The geometry / solid angle of detection
- The collimation
- The scatter
- The attenuation
- Detector response
- Detection efficiency of the detector
- Positron range, acolinearity, etc...

Improving the physical model enables:

Better quantification results

Better spatial resolution

...

Model measuring:

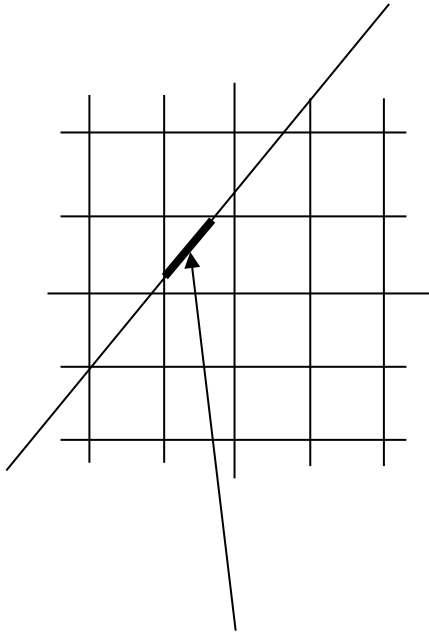
No approximation in the analytical calculation

Long time acquisition

Storage

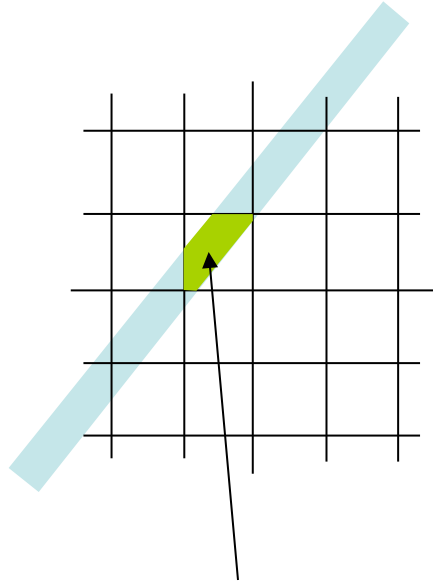
...

# -2- Integration line



$a_{ij}$  = intersection length

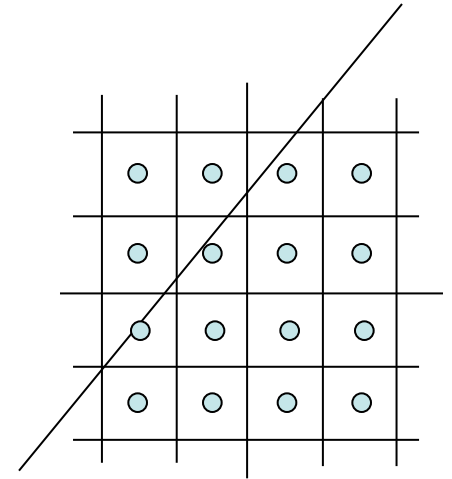
$$s_i(\vec{x}) = \delta(\vec{k}_i \cdot \vec{x} - \tau_i)$$



$a_{ij}$  = area of intersection

$$s_i(\vec{x}) = \text{rect}\left(\frac{\vec{k}_i \cdot \vec{x} - \tau_i}{w}\right)$$

Detector size



$a_{ij}$  = interpolation

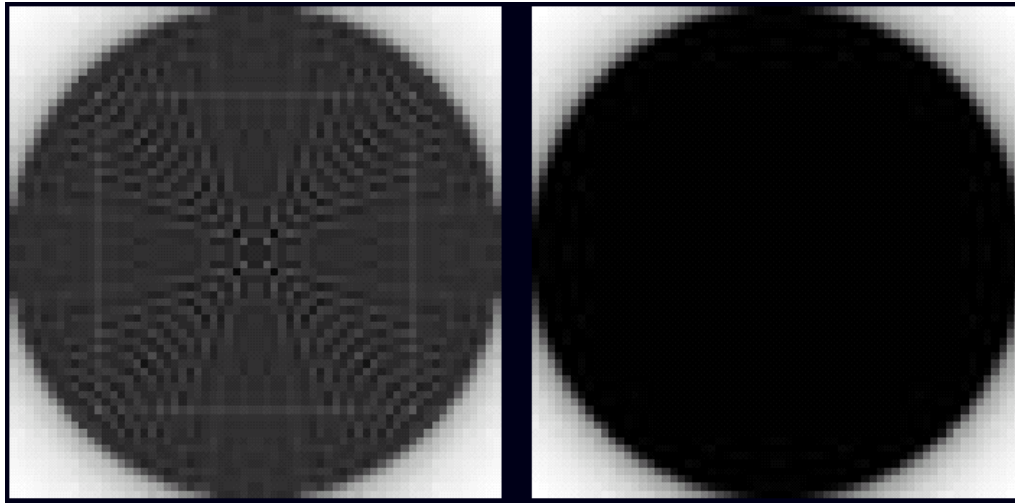
$$s_i(\vec{x}) = \text{tri}\left(\frac{\vec{k}_i \cdot \vec{x} - \tau_i}{\Delta r}\right)$$

Distance between lines

## -2- Examples...

Uniforme Sinogram

Back projection

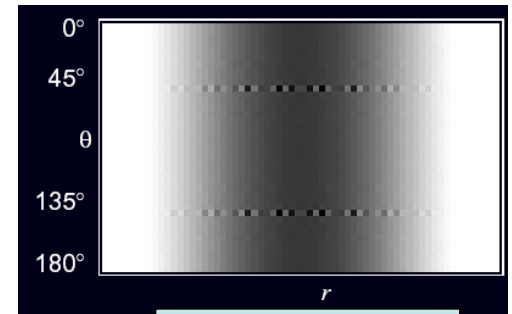


line

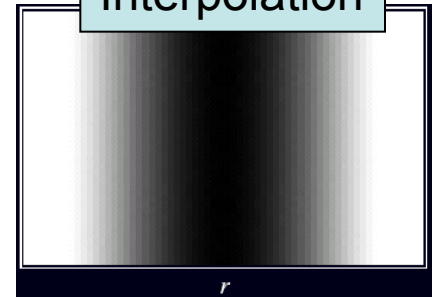
Area

Uniform object

Projection

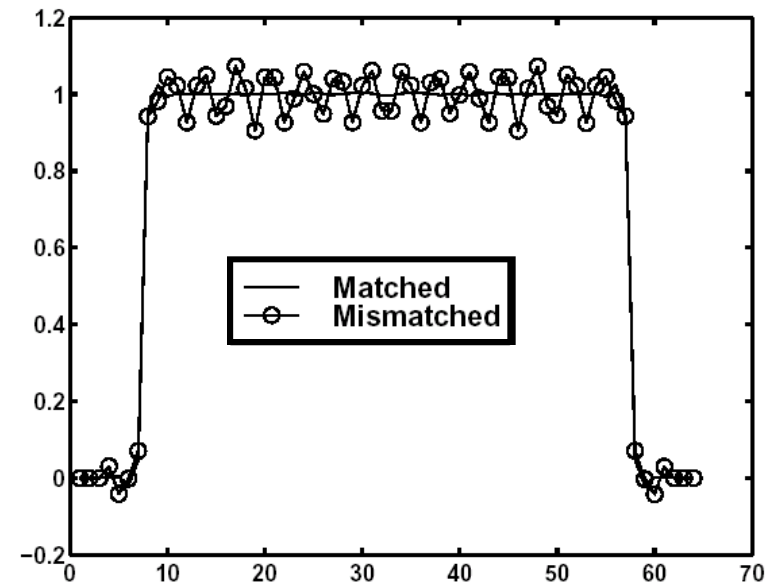
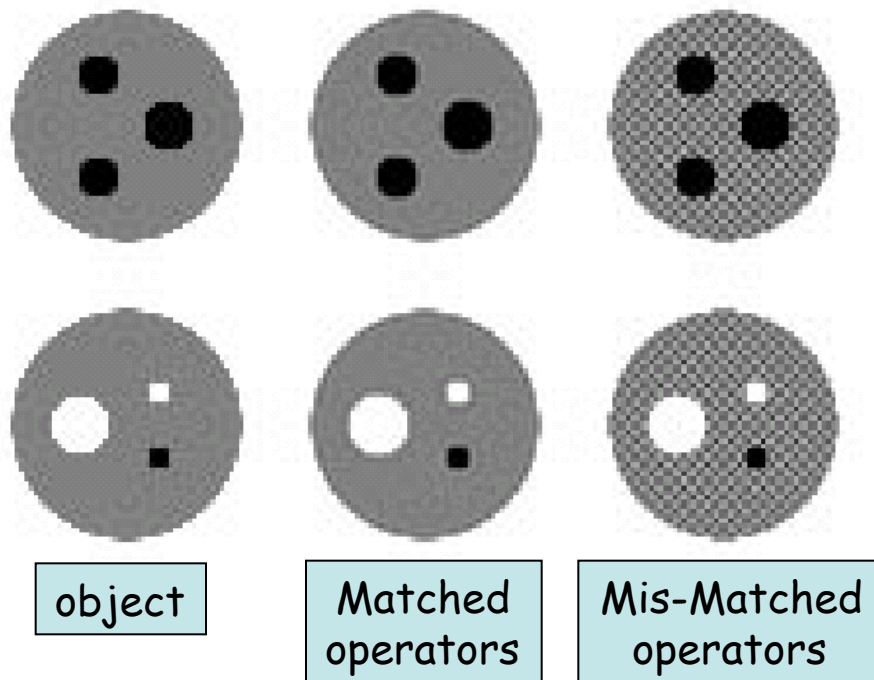


Interpolation



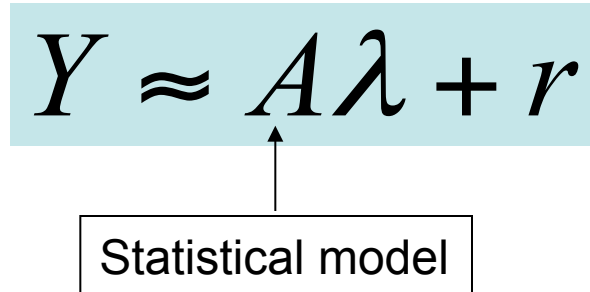
Area

# -2- Matched/Mismatched Projector/BackProjector operators





# -3- Statistical mode of measurement

$$Y \approx A\lambda + r$$


Statistical model

- Good model:
  - Variance reduction in image
  - Increasing computing time
  - Algorithm complexity
- Incorrect model
  - Statistics (dead time)
  - Model (transmission log)

# -3- Choice of the Statistical model

- No model:  $Y - r = A\lambda$       Resolve algebraically in order to find  $\lambda$ .

- Uniform gaussian noise: Least squares method, minimize  $\|Y - A\lambda\|^2$

- Not uniform gaussian noise: Weighted least squares method, minimize

$$\|Y - A\lambda\|_w^2 = \sum_{i=1}^{n_d} w_i (y_i - [A\lambda]_i)^2, \quad [A\lambda] \equiv \sum_{j=1}^{n_p} a_{ij} \lambda_j$$

- Poisson Model:  $Y_i \sim \text{Poisson}\{[A\lambda]_i + r_i\}$

# -4 & 5- Optimized criteria and used algorithm

Example:

Most used method: ML-EM

« *Maximum Likelihood* – *Expectation Maximisation* »

Optimization criteria

Algorithm

# Two steps per iteration

- 1<sup>st</sup> Step: E (« Expectation »)
  - Calculate the likelihood expectation
- 2<sup>nd</sup> Step: M (« Maximisation »)
  - Maximise the expectation.

# Mathematics derivation

## Definition:

$$\lambda_j$$

Mean number of desintegration in pixel  $j$

$$a_{ij}$$

Probability that a photon emitted in pixel  $j$  is detected in bin  $i$

$$a_{ij}\lambda_j$$

Mean number of photons emitted from pixel  $j$  and detected in bin  $i$

$$\bar{g}_i = \sum_{j=1}^m a_{ij} \bar{f}_j$$

Mean number of photons detected in bin  $i$

# Mathematics derivation

We have proved that  $g_i$  is a variable which statistics follows the Poisson law



The probability of detecting  $g_i$  photons is:

$$P(g_i) = \frac{e^{-\bar{g}_i} \bar{g}_i^{g_i}}{g_i!}$$

Example: Probability of detecting 5 with 3 as mean number of events

$$P(5) = \frac{e^{-3} 3^5}{5!} \approx 0.101$$

# Mathematics derivation

- Hypothesis on acquired data
  - The variables  $i$  are independents

$$P(g|\lambda)$$

Probability of observing the vector  $g$   
when the emission vector is  $\lambda$

$$=$$

$$\prod_{i=1}^n P(g_i)$$

Product of the individual probabilities

$$=$$

$$L(\lambda)$$

Likelihood function

# Mathematics derivation

- Find the maximum value  $\Rightarrow L(\lambda)$
- Calculate its derivative
- In order to maximize the likelihood, we use the following algorithm

$$l(\lambda) = \ln(L(\lambda))$$

$$l(\lambda) = \ln\left(\prod_{i=1}^n \frac{e^{-\bar{g}_i} \bar{g}_i^{g_i}}{g_i!}\right) = \sum_{i=1}^n \ln\left(\frac{e^{-\bar{g}_i} \bar{g}_i^{g_i}}{g_i!}\right) = \sum_{i=1}^n (-\bar{g}_i + g_i \ln(\bar{g}_i) - \ln(g_i!))$$



# Mathematics derivation

$$l(\lambda) = \sum_{i=1}^n \left( -\bar{g}_i + g_i \ln(\bar{g}_i) - \ln(g_i!) \right) \quad \text{avec} \quad \bar{g}_i = \sum_{j=1}^m a_{ij} \lambda_j$$

$$l(\lambda) = \sum_{i=1}^n \left( - \sum_{j=1}^m a_{ij} \lambda_j + g_i \ln \left( \sum_{j=1}^m a_{ij} \lambda_j \right) - \ln(g_i!) \right)$$

⇒ Probability to observe a projection from a mean image.

We want the image with the maximum probability of having  $g$

In other words, the vector  $\lambda$  for which  $l(\lambda)$  is maximum and considered as the best estimation of the solution.

# Mathematics derivation

- It was proved that  $l(\lambda)$  has a unique maximum.

- $\frac{\partial l(\lambda)}{\partial \lambda_j} = 0 \Rightarrow \text{maximum}$

$$\frac{\partial l(\lambda)}{\partial \lambda_j} = -\sum_{i=1}^n a_{ij} + \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}} a_{ij} = 0$$

$$l(\lambda) = \sum_{i=1}^n \left( -\sum_{j=1}^m a_{ij} \lambda_j + g_i \ln \left( \sum_{j=1}^m a_{ij} \lambda_j \right) - \ln(g_i!) \right)$$

# Mathematics derivation

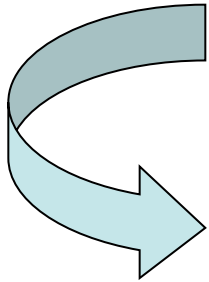
$$\frac{\partial l(\lambda)}{\partial \lambda_j} = -\sum_{i=1}^n a_{ij} + \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}} a_{ij} = 0$$

$$\boxed{\lambda_j} \frac{\partial l(\bar{f})}{\partial \bar{f}_j} = -\boxed{\lambda_j} \sum_{i=1}^n a_{ij} + \boxed{\lambda_j} \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}} a_{ij} = 0$$

$$\lambda_j = \frac{\lambda_j}{\sum_{i=1}^n a_{ij}} \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}} a_{ij}$$

# Iterative form

$$\lambda_j = \frac{\lambda_j}{\sum_{i=1}^n a_{ij}} \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}} a_{ij}$$



$$\lambda_j^{(k+1)} = \frac{\lambda_j^{(k)}}{\sum_{i=1}^n a_{ij}} \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}^{(k)}} a_{ij}$$

# Description

Measured Projection

$$\lambda_j^{(k+1)} = \frac{\lambda_j^{(k)}}{\sum_{i=1}^n a_{ij}} \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}^{(k)}} a_{ij}$$

Normalization factor

Estimated projection

Image<sup>(k+1)</sup> = Image<sup>(k)</sup> x back projection normalized with  $\frac{\text{Measured projection}}{\text{Estimated projection}}$

# ML-EM Algorithm

- Multiplicative method
- Positive or null solution
  - Initial values at 0 remain 0
  - Positive initial value remain positive
- Conservation of the global activity in the image
- Slow convergence
- For a low number of iterations
  - Cold zones: excellent reconstruction
  - Hot zones: reconstruction  $<$  FBP
- For a high number of iterations
  - Cold zones: excellent reconstruction
  - Hot zones: Noisy images (bias near to 0)

# Noise at Convergence

- At convergence,
  - « Perfect » reconstruction of the counts number in each pixel
- However,
  - No correlation between neighbouring pixels.
  - High Poisson noise level  $\Rightarrow$  Chessboard effect
- Corrections
  - Stop the iterations (need to define a stop criteria...)
  - Penalization function

# Research domains

- Convergence acceleration
- Problem regularization
  - Penalty function
  - Introduction of an A PRIORI knowledge



# Convergence acceleration

- Algorithm OS-EM

$$\lambda_j^{(k+1)} = \frac{\lambda_j^{(k)}}{\sum_{i=1}^n a_{ij}} \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}^{(k)}} a_{ij}$$

OS-EM = ML-EM applied  
on a subset  $S$   
 $S=1 \Rightarrow$  ML-EM

Convergence has not been  
proved but seems to be  
similar to that of ML-EM.

$$\lambda_j^{(k+1)} = \frac{\lambda_j^{(k)}}{\sum_{i \in S} a_{ij}} \sum_{i \in S} \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}^{(k)}} a_{ij}$$

Acceleration factor  $\sim S$

Adequate choice of subsets

# Regularization

- Criteria:
  - Estimate projection  $\sim$  measured projection.
- Replaced by:
  - (a) Estimated projection  $\sim$  measured projection.
  - (b) Low noise obtained image.
- The introduction of an a priori knowledge on the image = regularization
  - Promote convergence!



Find  $\lambda$  to maximize (a) and (b)

# Mathematics derivation

Bayes theorem:

Probability of observing the vector  $g$   
when the emission vector is  $f$

A priori knowledge on the image

$$P(\lambda|g) = \frac{P(g|\lambda)P(\lambda)}{P(g)}$$

A posteriori probability

A priori knowledge on the projections

# Algorithm MAP: Maximum a Posteriori

Consider the logarithm:

$$P(\lambda|g) = \frac{P(g|\lambda)P(\lambda)}{P(g)} \Rightarrow \ln P(\lambda|g) = \underbrace{\ln P(\lambda|g)}_{\text{A posteriori probability}} + \underbrace{\ln P(g|\lambda)}_{\text{Likelihood}} + \underbrace{\ln P(\lambda)}_{\substack{\text{a priori} \\ \downarrow \\ \text{« prior »}}} - \underbrace{\ln P(g)}_{\text{Constant}}$$

MAP = ML if no a priori information  $\Rightarrow$  ML = special case of MAP

MAP = ML penalized, the penalty being the a priori knowledge.

# A priori example

Gibbs a priori  $\Rightarrow$  local image smoothing

$$P(\lambda) = C e^{-\beta U(\lambda)}$$

$U$  : Energy function of  $\lambda$

$\beta$  : A priori weighting

$C$  : Normalization constant

$$\ln P(\lambda|g) = \sum_{i=1}^n \left( - \sum_{j=1}^m a_{ij} \lambda_j + g_i \ln \left( \sum_{j=1}^m a_{ij} \lambda_j \right) - \ln(g_i!) \right) - \beta U(\lambda) + K$$

$$K = \ln C - \ln P(g) \quad : \text{Constant independent of } \lambda$$

# Maximize likelihood

- Derive the likelihood in order to maximize  $\lambda$

$$\frac{\partial P(\lambda|g)}{\partial \lambda_j} = -\sum_{i=1}^n a_{ij} + \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}} a_{ij} - \beta \frac{\partial}{\partial \lambda_j} U(\lambda_j) = 0$$

$$\lambda_j^{(k+1)} = \frac{\lambda_j^{(k)}}{\sum_{i=1}^n a_{ij} + \beta \frac{\partial}{\partial f_j} U(\lambda_j^{(k)})} \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}^{(k)}} a_{ij}$$

# Example of function $U$

- A quadratic a priori:

$$\frac{\partial}{\partial \lambda_j^{(k)}} U(\lambda_j^{(k)}) = \sum_{b \in N_j} w_{jb} (\lambda_j^{(k)} - \lambda_b^{(k)})$$

$N_j$  : set of points neighbouring pixel  $j$

Si  $b \sim j \Rightarrow$  the term is zero  $\Rightarrow \lambda^{(k+1)}$  same to ML-EM

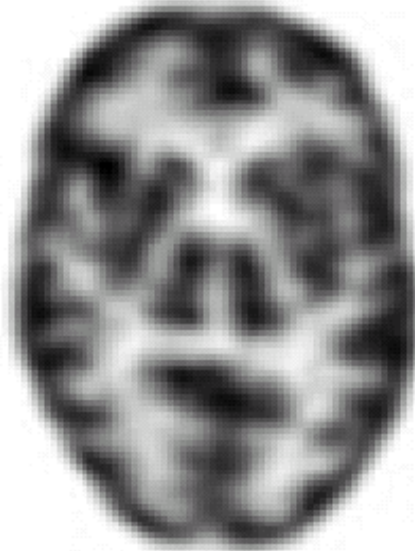
Si  $j > b \Rightarrow$  the term  $>0 \Rightarrow \lambda^{(k+1)} <$  a ML-EM

Si  $j < b \Rightarrow$  the term is  $<0 \Rightarrow \lambda^{(k+1)} >$  a ML-EM

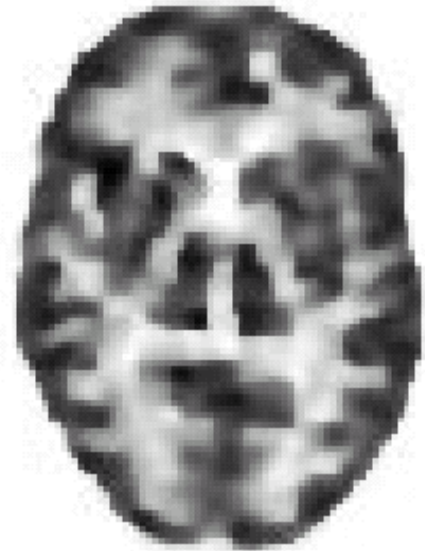
# Examples



object



quadratic a priori



Hubert a priori



# Some remarks

$$\lambda_j^{(k+1)} = \frac{\lambda_j^{(k)}}{\sum_{i=1}^n a_{ij} + \beta \frac{\partial}{\partial f_j} U(\lambda_j^{(k)})} \sum_{i=1}^n \frac{g_i}{\sum_{j'=1}^m a_{ij'} \lambda_{j'}^{(k)}} a_{ij}$$

Possibility of negativity

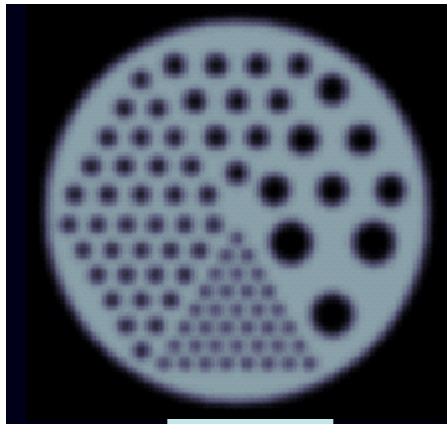
⇒ Keep a low value of  $\beta$  so that values remain positive

The a priori smooths also the edges ⇒ Loss of resolution

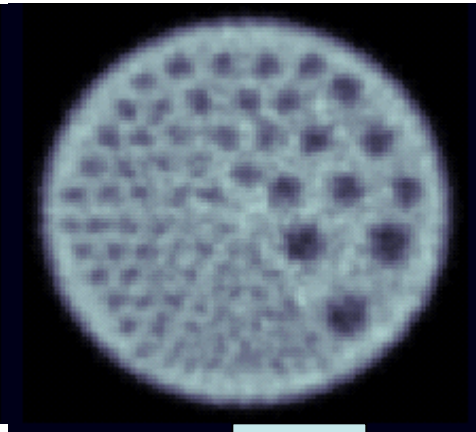


Modify the a priori  
Introduce anatomical information

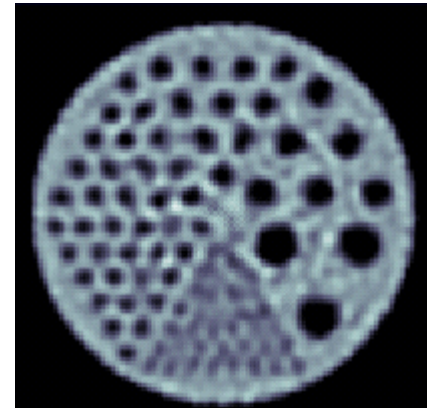
# Illustrations



object

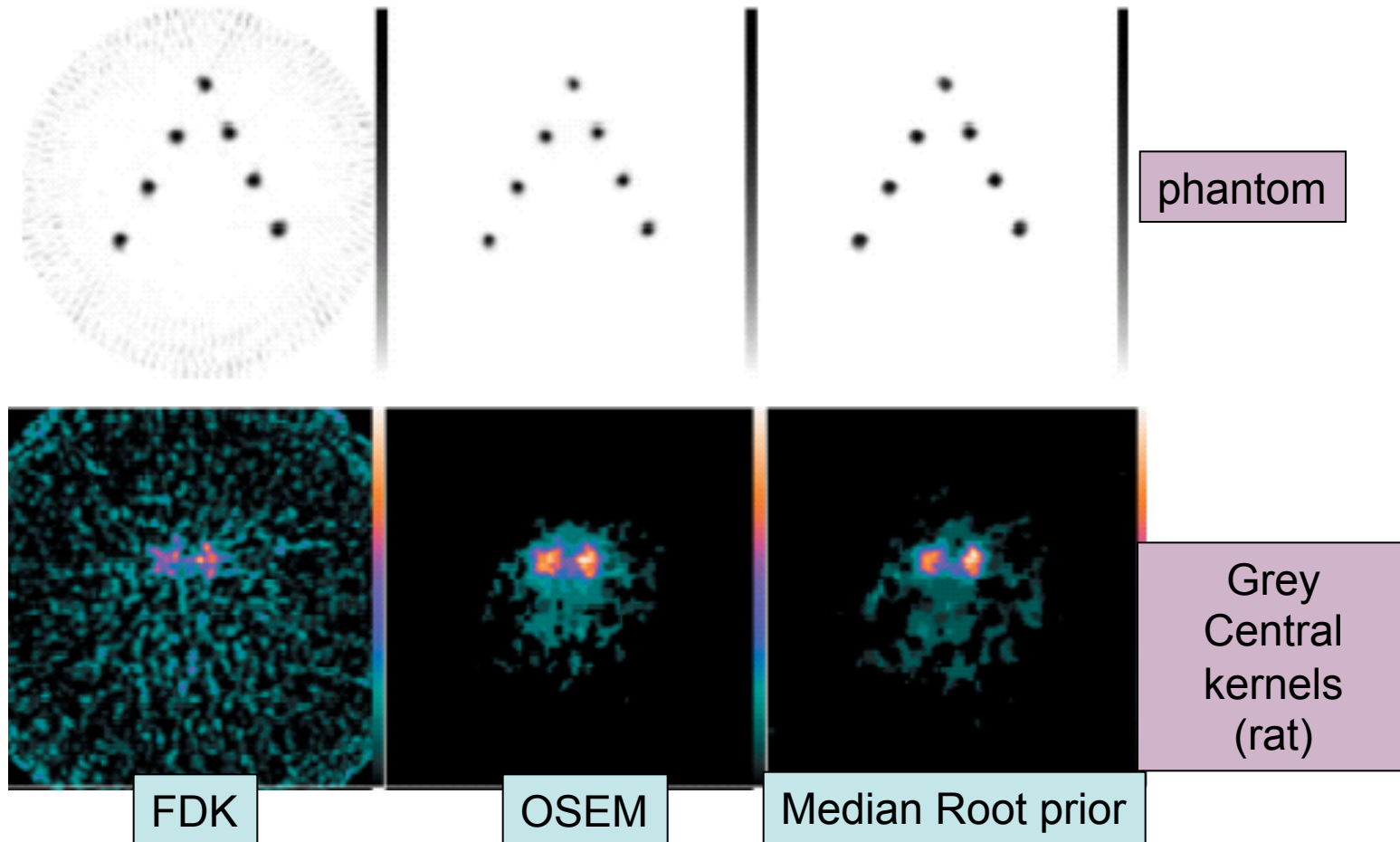


FBP



Iterative + regularisation

# Illustrations



# References

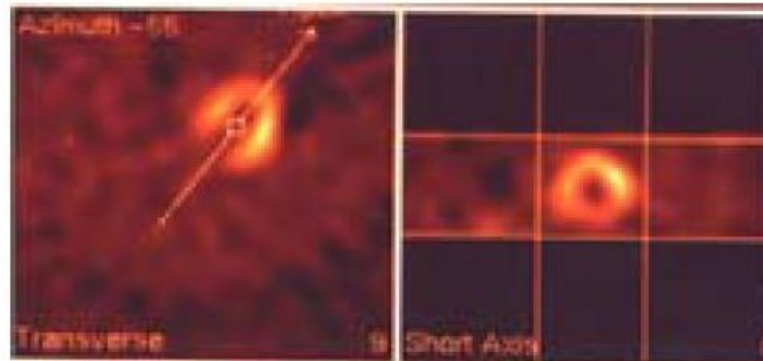
- F. Beekman, *Discrete Reconstruction Methods*, NSS-MIC 2000.
- J. A. Fessler, *Statistical Method for Image Reconstruction*, NSS-MIC 2001.
- P.P. Bruyant, *Analytic and Iterative Reconstruction Algorithms in SPECT*, JNM 2002.

# Type of tomographic studies

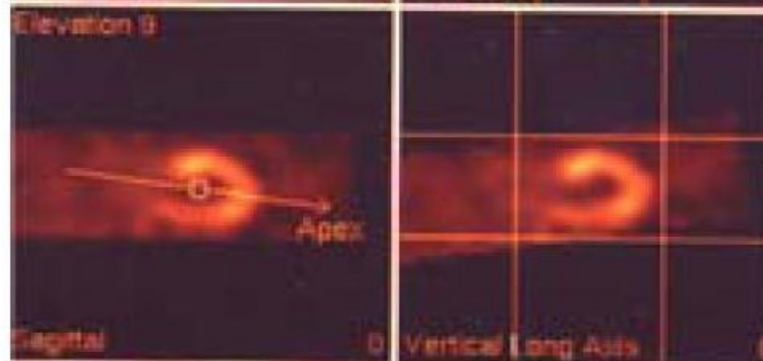
- Heart
  - Myocardec (thallium, MIBI...)
  - Ventricular cavities
  - Gated SPECT
- Brain
- Lungs
- Bone
- Others (peptides, antibodies...)

# Slices plans in myocardial imaging

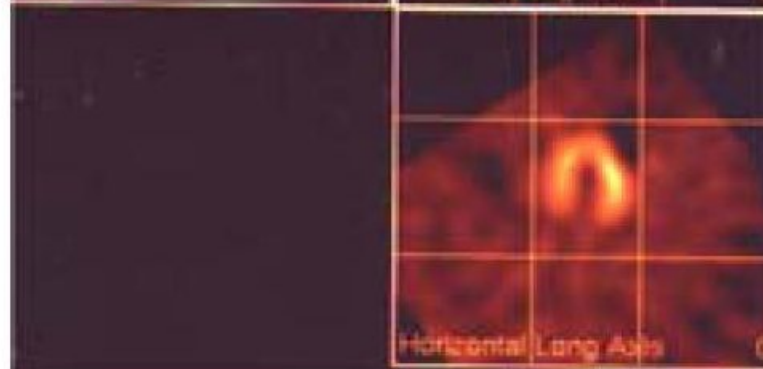
Small axis



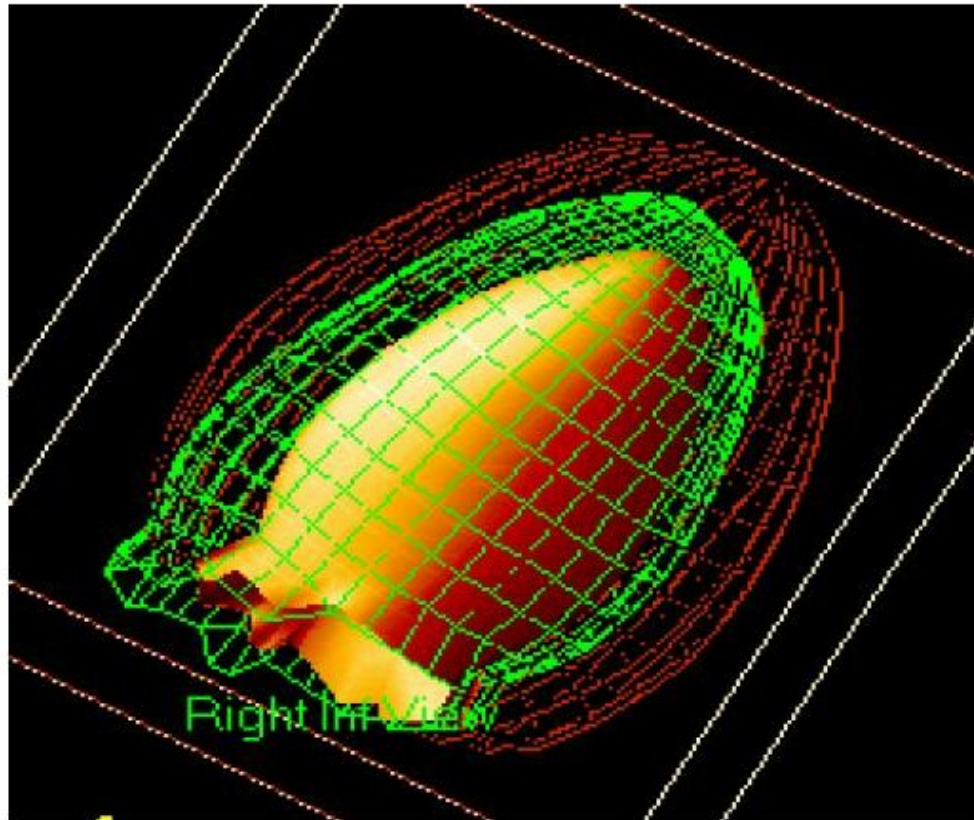
Big axis



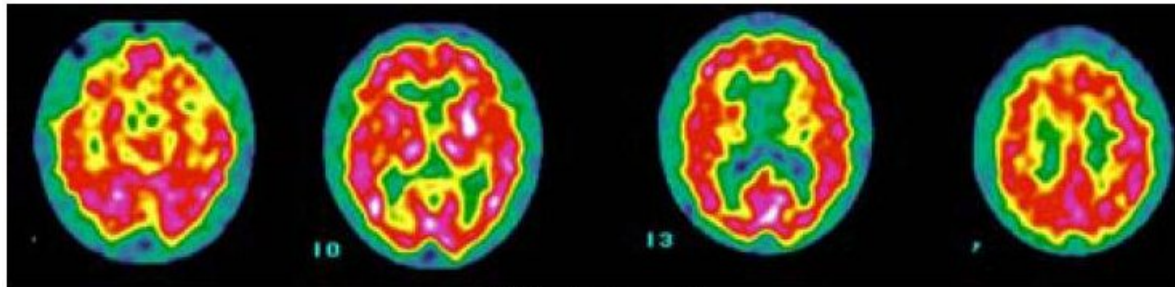
Horizontal



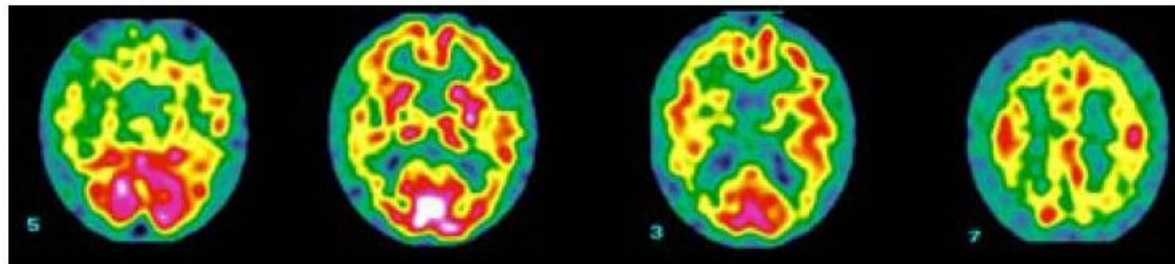
# Representation of 3D contours



# Cerebral tomography



Sujet normal



Maladie d'Alzheimer



# Endocrine tumor in the liver

